

Party bans: deterrence or backlash?

Evidence from the Basque Country

Andreu Arenas*

November 20, 2019

Abstract

This paper investigates the electoral effects of party bans, studying the case of the ban on Batasuna, the political wing of ETA. Exploiting the finite and heterogeneous length of the ban across municipalities, I find that a longer ban has a negative effect on electoral support for the targeted party. This effect is explained by the extent of the immediate loss in support in treated municipalities under the ban, observable in this instance because Batasuna called for a null vote. This pattern and further heterogeneous effects are consistent with voters learning from the new electoral scenario rather than with a direct cost attributable to the party no longer having institutional representation. This suggests that party bans may be used by incumbents to reduce their challengers' strength by triggering an electoral reshuffling, with voters learning about new parties and potentially switching allegiances.

*University of Barcelona and IEB. andreu.arenas@ub.edu. I am grateful to the editors, two anonymous reviewers, Cesc Amat, Manuel Bagues, Claude Berrebi, Bruce Bueno de Mesquita, Micael Castanheira, Michael Dorsch, Tiberiu Dragu, Ruben Durante, Ruben Enikolopov, Raphaël Franck, Javier Gardeazabal, Amuitz Garmendia, Eric Gould, Andrea Ichino, Esteban Klor, Andrew Long, Clément Malgouyres, François Maniquet, Andrea Mattozzi, Katja Michaelowa, Luis Miller, Jordi Muñoz, Maria Petrova, Pablo Querubín, Guillem Rico, Arturas Rozenas, Cyrus Samii, David Yanagizawa-Drott, Asaf Zussman, and seminar and conference participants for helpful comments. This paper is based on the third chapter of my PhD dissertation at the EUI. The hospitality of Aalto University, the Hebrew University of Jerusalem, and NYU, where part of this work was carried out, is gratefully acknowledged.

Many democratic countries restrict the participation of political parties whose behavior or viewpoints are deemed to go beyond what is tolerable. In recent decades, over 50 political parties have been banned across Europe, including parties in Spain, Germany, Turkey, the United Kingdom, and Bulgaria. In 2013, the far right extremist Golden Dawn came third in the 2013 Greek election, after EU officials unsuccessfully called for its ban.¹ In Africa, the number of bans on ethnic, religious and regional parties has increased in sub-Saharan countries since the 1990's.²

The variation in party ban practices across countries has led to a various studies that discuss the normative and legal arguments surrounding party bans. Fox and Nolte (1995), Issacharoff (2007) and Capoccia (2013) argue that there is an inevitable democratic dilemma in party bans, consisting of a tension between procedural and substantive conceptions of democracy: democratic elections may require, as a precondition to the right of participation, a commitment to the preservation of the democratic process, but at the same time, this limits the scope of democratic deliberation, calling into question the legitimacy of the political process. A common concern in these studies is striking the right balance between the danger of extremist parties gaining excessive power and implementing anti-democratic policies, and the danger that party bans might be used to insulate incumbents from electoral challenge. In this spirit, according to the Guidelines on the Prohibition and Dissolution of Political Parties and Analogous Measures issued by the European Commission for Democracy through Law (Council of Europe - Venice Commission, 1999), party bans are an exceptional measure that requires sufficient evidence that a party is a real threat to the constitutional order or citizens' fundamental rights and freedoms.

Discussions as to the desirability of party bans often include instrumental arguments based on the expected electoral effects of such bans. However, besides a short-run mechanical

¹Financial Times, April 2013.

²Other examples include the Islamists Welfare Party (Turkey) and the Muslim Brotherhood's Freedom and Justice Party (Egypt), the far-right secessionist Vlaams Blok (Belgium), the far-right Kach (Israel), the Communist Party of Lithuania, the Communist Party of Germany (PKD), the Macedonian secessionist Ilinden—Pirin (Bulgaria), the left-wing, independentist Sinn Féin and Republican Clubs (UK) and Batasuna (Spain). Casal Bértoa and Bourne (2016) provide a comprehensive list of party bans in Europe since 1945. In Argentina, all Peronist entities were banned from elections after 1955. Party bans are pervasive in sub-Saharan Africa (Basedau and Moroff, 2011). Besides party bans in democracies, partially or nominally democratic institutions such as free elections with constraints on the opposition are increasingly being used by autocrats to remain in power (Gandhi and Lust-Okar, 2009).

exclusion, it remains an open empirical question whether party bans increase or reduce support for the targeted parties. According to Capoccia (2005), declaring an extremist party illegal might deprive the extremist movement of organizational resources and social visibility, and prevent extremists from influencing policy in the short-run. However, party bans carry the risk of significant electoral costs. First, such measures might encounter strong principled opposition in public opinion because they contradict democratic principles too starkly; second, they might force the extremists to go underground and adopt other, possibly violent, tactics, which pose a different kind of problem to the incumbents. Bale (2007) compares party bans in Belgium, Spain, and Turkey, and concludes that they are not associated with an increased popularity of the targeted parties. Casal Bértoa and Bourne (2016) compare party bans in Germany, Spain, and Turkey, and conclude that party bans are associated with periods of electoral volatility. Minkenberg (2006) reports that the German fascist Sozialistische Reichspartei, which was banned in 1952 but which re-emerged in 1969, did not manage to clear the 5% electoral threshold, suggesting that the ban reduced party support. The Czech Republic banned the marginal, far-right Workers' Party in 2010, but a successor party later obtained a similar level of electoral support (Downs, 2012).³ All these qualitative studies suggest that party bans are unlikely to lead to an electoral backlash; nevertheless, it is difficult to know whether these aggregate outcomes are indeed caused by party bans, or if instead party bans are simply more likely to occur in periods when the targeted parties are expected to be weak, or in contexts of political instability. This limits our ability to understand the reasons why party bans might be used in the first place as well as the reasons why party bans may have an effect on electoral outcomes.

This paper contributes to this literature by examining a case study that allows me to construct a counterfactual to disentangle the effect of the party ban from confounding effects. This case study is the ban on Batasuna, a leftist-independentist political party in the Spanish region of the Basque Country that was outlawed because of its connection with the Basque terrorist organization ETA. The ban on Batasuna had a finite and heterogeneous length across municipalities: in the first electoral term, the party was banned in all municipalities, while in the second, it was banned only in about half of them, and in a third electoral term,

³Basedau and Moroff (2011) study 12 party bans in sub-Saharan Africa, with mixed findings.

the ban on the party was lifted. Crucially, these two groups of municipalities had reacted very similarly to aggregate popularity shocks before the heterogeneous ban (i.e., they had been following parallel trends in electoral outcomes). This allows me to empirically examine the effects of party bans by comparing the outcomes of two groups of similar municipalities that were exposed to bans of different intensity. Furthermore, the ban on Batasuna is an interesting case study because the party had a high level of pre-ban support, above 20% across municipalities; and because the ban was promoted by the Spanish Parliament against the will of the Basque Parliament.

The main result of the paper is that a longer ban had a negative effect on support for Batasuna. The effect was large and persistent for at least two elections following the lifting of the ban. Furthermore, using data on episodes of street violence perpetrated by Batasuna supporters, the results indicate that the extension of the ban for an additional electoral term led to a one-month spike in street violence in treated municipalities, but it had no effects during the following year. The interpretation of the results as a causal effect of the ban relies on the assumption that both groups of municipalities would have followed parallel trends in outcomes in the absence of the heterogeneous length of the ban. An inspection of pre-treatment trends suggests that this is a plausible assumption.

A party ban is a repressive measure that prevents a specific political party from contesting elections and obtaining any of the benefits from being in office or holding seats in parliament. Hence, previous studies of the effects of party bans mainly consider two sets of possible mechanisms. A first set of mechanisms concerns voters' reactions to the ban as a policy. These are assumed to be negative because of the democratic contradiction that it might entail, possibly generating sympathies towards the party among voters who are simply opposed to the ban. For instance, Van Spanje and De Vreese (2015) use data from a representative sample of Dutch voters interviewed before and re-interviewed after the unexpected court decision to prosecute the far-right MP Geert Wilders for hate speech, and find that it substantially enhanced his party's appeal. In the context of counter-terrorist policies, a common thread in the literature is that more indiscriminate actions are more likely to result in a backlash (Dragu and Polborn, 2014; Benmelech et al., 2015; Dragu, 2017). Likewise, the empirical literature on repression, which has focused on the effect of state violence, has shown that

indiscriminate violence systematically reduces political support for the perpetrator, both at the individual level and overall in the most affected areas (Balcells, 2012; Rozenas et al., 2017; Lupu and Peisakhin, 2017). The negative effect of the ban on support for Batasuna, however, suggests that this set of mechanisms did not play a major role in this case.

A second set of mechanisms concerns the relative disadvantages suffered by the targeted party because of its inability to contest elections and hold seats in city councils. This might be associated with a loss of support for two reasons: the direct material disadvantages of not holding parliamentary representation and its effect on the party’s mobilization capacity, and voters’ increased information about competing politicians that would not have been elected otherwise. For instance, Dinas et al. (2015) show that small parties that barely cross the electoral threshold to enter parliament substantially improve their electoral performance in the following election, especially in new democracies, with weaker party branding.⁴

The second main contribution of this paper is to study the heterogeneous effects of the ban, with the aim of understanding the role of these mechanisms. First, I study how the electoral effects of the ban unfold over time. This is possible because when the ban did not apply to all municipalities, Batasuna continued to contest the elections in those places where it was banned by calling for a null vote. Because the number of null votes is typically very close to zero, the immediate loss in support for the party under the ban is observable. If the direct costs of not holding city council seats are the main mechanism, this should not be detectable immediately, but rather emerge gradually over time. On the other hand, if the main mechanism is voter learning about new parties, one would expect a significant correlation between the very immediate effect of the ban on support for Batasuna in 2007 and the effect once the ban was lifted. First, this is because wherever Batasuna voters are more likely to abandon the party under the ban, the composition of the remaining parties in the city council is less likely to be the same, because not all parties gain the same from the ban: I find that those that gain the most are small local parties, with less known platforms. Second, because under the ban, when voting for Batasuna does not translate into seats, the party’s lack of potential for policy influence from the outset increases the incentives for instrumental

⁴The related literature on campaign advertising finds significant effects on electoral outcomes, especially of non-dominant parties (Larreguy et al., 2018; Spenkuch and Toniatti, 2018; Da Silveira and De Mello, 2011)

voters to learn about new parties. These mechanisms would lead to a positive correlation between the immediate and the long-run loss in support due to the ban.

The results indicate that the very immediate loss of electoral support in treated municipalities predicts very well the long-run loss of support following the lifting of the ban, which is consistent with voter learning about new candidates being the main disadvantage of the targeted party due to the ban, rather than the direct cost from being excluded from the city council. This suggests that party bans may be used by incumbents to reduce their challengers' strength by triggering an electoral reshuffling, with voters learning from the new scenario and potentially switching parties.

I also examine in which municipalities Batasuna is more or less likely to lose votes. I find that the vote loss is especially pronounced in municipalities with a higher number of ETA attacks in the past, which is consistent with the idea that more indiscriminate policies are perceived as illegitimate and are more likely to backfire, even if overall this seems to be of secondary importance because of the negative effect of the ban on support for Batasuna. I also find that the effect is larger in municipalities where other nationalist parties and federal parties had similar vote shares in the local elections preceding the 2007 ban. This also suggests that the effect of the ban might depend on the incentives for party switching, and that the existence of alternative organizations in competition with the targeted party might be important, as proposed by Capoccia (2005). In both cases, the immediate and the long-run effects show a similar pattern, indicating again a persistent rather than a gradual pattern in the vote loss. Finally, I also find that the effect of the ban is greater wherever the pre-ban variance of the electoral support for the party is greater (conditional on the pre-ban level of support), which is also consistent with voters' propensity to switch parties playing a role in the effect of the ban. Nevertheless, because these municipalities could differ in other dimensions, the interpretation of these heterogeneous effects is suggestive, and further investigation of these hypotheses and mechanisms is left for future research.

The ban on Batasuna

The Basque Country is a region in Northern Spain with a multi dimensional party system. Every party is characterized by a policy position on the left-right spectrum and in relation to its stance regarding regional autonomy, with positions on the latter ranging from full state centralization to Basque independence. Apart from their policy positions, the electorate is similarly divided by cultural identity, which correlates with a range of preferences regarding decentralization. The main parties contesting the Basque elections can be divided into Federal parties (with candidatures everywhere in Spain) and Basque parties (with candidatures only in the Basque Country, typically with a preference for more regional autonomy). Among the latter, Batasuna was the main leftist and independentist party in the region until its ban in 2003. Over the period of analysis, this party adopted different names and electoral brands, including “Herri Batasuna”, “Euskal Herritarrok”, and “Acción Nacionalista Vasca”. For simplicity, I refer to these parties and to the political representation of the “izquierda abertzale” (leftist-independentism) as Batasuna. Before its ban, Batasuna used to obtain between 20 and 25% of the vote in the Basque local elections.

Politics in the Basque Country have been heavily influenced by the existence of Euskadi Ta Askatasuna (ETA), a terrorist organization in favor of the independence of the region.⁵ ETA was created in 1958 (during Franco’s dictatorship in Spain) and was active until 2011, when it announced a permanent end to its armed activities.⁶ Over this period, ETA killed more than 800 people, mostly between the 1970’s and the 1980’s. Besides its direct victims, ETA’s terrorism has had important economic consequences for the region and has also influenced Basque politics. One of the most politically relevant consequences of the existence of ETA has been the ban on Batasuna, which used to represent its ideological space (leftist-independentism) and never rejected ETA’s terrorism.

In June 2002, the Spanish Parliament passed a new Law of Political Parties, with the

⁵A number of contributions have studied other questions related to identity, political conflict and terrorism in the Basque Country (de la Calle, 2007; Barros et al., 2009; Gil-Alaña and Barros, 2010; Gardeazabal, 2011; de la Calle and Sánchez-Cuenca, 2013; Ansolabehere and Puy, 2016), and its economic effects (Abadie and Gardeazabal, 2003)

⁶During this period, ETA held a number of ceasefires. The last ceasefire started in September 2010. In January 2011, ETA announced that this ceasefire would be permanent and verifiable by international observers. Eventually, on October 2011, ETA announced a definitive cessation of all its armed activities.

support of more than 90% of its members, with the following aim: “... to guarantee the operation of the democratic system and the citizens’ essential freedoms, avoiding the possibility that a political party could, in a reiterative and serious fashion, threaten this democratic regime of freedom, justify racism and xenophobia or lend political support to violence and the activities of terrorist groups...”⁷

A few weeks later, the Council of Ministers asked the Supreme Court of Spain to ban Batasuna. In March 2003, following a period of deliberation, the Supreme Court banned Batasuna owing to its links with ETA and its failure to reject terrorism.⁸ The ban was not the result of bargaining among the existing political agents in the Basque Country, which mostly opposed it; rather, it was a process led by the Spanish Parliament and subsequently enforced by the Courts. The Law of Political Parties was passed by the Spanish Parliament, with a majority government under the People’s Party (PP), the party that had suffered the most victims of ETA terrorism. Indeed, this was a period of increasing international concern about terrorism, especially following the 9/11 terror attacks.⁹

The political context surrounding the 2003 ban on Batasuna was a general non-conciliatory stance by the Spanish government and judiciary towards Basque political organizations and the political majorities in the Basque parliament. For instance, in February 2003 the Spanish National Court closed down *Egunkaria*, the only fully Basque language newspaper in circulation due to allegations of links with ETA, allegations that would be withdrawn seven years later. In June 2003, the Spanish parliament passed a law extending the maximum length of effective prison sentences from 30 to 40 years. A majority of members of the Basque regional parliament voted against the law and accused the Spanish government of basing its policies on a desire for revenge. And in November 2003, the Spanish government passed a new law introducing prison sentences of between 3 and 5 years for anyone calling for a referendum on autonomy, in response to the development of a new proposal on autonomy made by the Basque government (a coalition of moderate Basque nationalists and federal leftists).

⁷Organic Law 6/2002 of Political Parties. Voted in favor: PP, PSOE, CiU, CC, PA. Voted against: PNV, EA, BNG, ERC, ICV, CHA.

⁸The political organizations that were outlawed by that judicial sentence included Herri Batasuna, Euskal Herritarrok and Batasuna.

⁹See Horz (2016) for a game-theory model of the determinants of leaders’ incentives to ban parties.

After its ban, Batasuna adopted different strategies to maintain its electoral presence. In 2003, in the first local election after the ban, they were completely unsuccessful in their efforts.¹⁰ With the aim of participating at the 2007 local election, Batasuna made use of two political parties, the first of these, “Abertzale Sozialisten Batasuna”, however was fully outlawed by the Courts. The second, “Eusko Abertzale Ekintza-Acción Nacionalista Vasca” (EAE-ANV), was an old party that had been inactive for decades. In this case, the party was revived by Batasuna very shortly before the elections, so that the Courts and the police had insufficient time to find enough evidence of links to the previously banned organizations to ban the party before the election. After the Public Prosecutor failed to prove that EAE-ANV was the same party as Batasuna, instead of giving up, he decided to denounce the municipalities for which he thought the evidence of a connection with Batasuna was the strongest. More specifically, the public prosecutor denounced the lists with a relatively large number of former Batasuna candidates (3 or more) and councilors (1 or more), or some other visible connection. The courts, in the main, followed the public prosecutor’s report and banned these electoral lists, which affected around 50% of the municipalities. Although Batasuna denied having any connection to EAE-ANV until the Courts’ final verdict was announced, shortly afterwards they endorsed EAE-ANV, revealing that it was part of their strategy to recover their positions in the institutions.¹¹ Even though Batasuna wanted to conceal the fact that they were behind EAE-ANV and that presenting its former candidates could raise suspicions, it should be borne in mind that the public prosecutor established his criteria ex-post (i.e., after the electoral lists had been presented), and that this was the only instance in which an *abertzale* organization - among the various organizations that were banned around that period - was partially banned based on local criteria.¹²

With the aim of participating at the 2011 local election, Batasuna once more created another party, which was meant to lead a larger candidature named “Bildu”, jointly with smaller parties such as EA (centrist independentists) and Alternatiba (a Basque splinter

¹⁰In 2003, they created a new party (“Autodeterminaziorako Bilgunea” - AuB), but the Supreme Court invalidated AuB’s lists owing to its links with the previously banned organizations.

¹¹Eventually, in 2008, EAE-ANV was also fully banned, but this did not occur until after the 2007 election when they won representation in many municipalities.

¹²For instance “Autodeterminaziorako Bilgunea” (2003) or “Abertzale Sozialisten Batasuna” (2007), and “Demokrazia Hiru Milioi” and “Askatasuna”, which were fully banned from the 2009 Basque regional election.

group from IU, federal leftists). For the first time, Batasuna decided to explicitly reject ETA's violence, with the aim of being allowed back into the political arena. Despite this, on the 1st May 2011, Spain's Supreme Court invalidated Bildu's electoral lists and forbade the inscription of the new Batasuna party on the registry of political parties, following police reports of links with ETA. However, a few days later, on May 5th, the Constitutional Court revoked the Supreme Court decision and allowed Bildu to contest the 2011 election to be held on May 22nd, in a close decision: with six judges voting in favor, five against (Garmendia Madariaga, 2011). In short, Batasuna was banned for two local elections (2003 and 2007), although the 2007 ban applied only to a subset of municipalities.

Effects on electoral support

Figure 1 shows the level of turnout at every local election in the Spanish region of the Basque Country between 1987 and 2015, the sample being divided between municipalities where Batasuna was banned in 2007 and municipalities where it escaped this ban.¹³ Turnout had followed parallel trends across the treated and control municipalities before the ban, and it did not undergo any substantial changes after the lifting of the ban (in 2011 and 2015).

Figure 2 reports null votes as a % of turnout (note that the y-axis scale changes across these first seven figures). Null votes were very close to zero before the ban. In 2003, when Batasuna was banned everywhere and the party called for a null vote, null votes as a percentage of turnout jumped from almost zero to more than 15%, representing around 70% of Batasuna's average pre-ban vote share. In 2007, Batasuna was deemed legal in 98 of the 251 Basque municipalities, and, hence, in those municipalities null votes returned to being close to zero. In the remaining municipalities, null votes remained at around 15%. Note that municipalities escaping the 2007 ban received more null votes in 2003. As the following figure shows, this was to be expected, because they also had more pre-ban support. After taking the pre-ban support difference into account, it seems that, if anything, the municipalities that escaped the 2007 ban dealt slightly worse with the 2003 ban. This outcome is examined in further detail in a later section dedicated to identification threats.

¹³Local elections take place every four years, in all municipalities at the same time.

Figure 3 reports the vote share of Batasuna over time. In 2003, when Batasuna was banned everywhere, it was zero. Before that, treatment and control municipalities had followed parallel trends. The pre-ban gap in support across treatment and control municipalities increased under the 2007 ban, precisely because of the ban, and persisted thereafter, suggesting that a longer ban led to relatively less post-ban support. Looking at figure 3 alone, however, it might be thought that in spite of these differences, overall the ban backfired and that whenever Batasuna returned to the electoral arena under regular conditions, it obtained more support than previously. However, the larger vote shares obtained by Bildu can be attributed to the presence of other political actors in this coalition, such as the centrist and independentist Eusko Alkartasuna (EA), which prior to the ban had contested the elections either alone or in coalition with the conservative nationalists.¹⁴ Figure 4 reports the sum of the votes for all nationalist parties, including Batasuna, EA, the conservative nationalist PNV, EA-PNV coalitions, Aralar, a moderate independentist party, and Null votes (capturing support for Batasuna in 2003 and 2007). The figure suggests that the 2007 ban led to a decline in support for Basque Nationalist parties in treated municipalities, a trend that persists after the lifting of the ban.

Figure 5 reports the vote share of Federal Leftist parties (mainly the Socialist Party, United Left, the Green Party, and Podemos). Although prior to the 2007 ban the vote share of these parties was slightly lower in treated municipalities, in 2007 it increased above the level of control municipalities, and it remained at a similar higher level after the lifting of the ban. In contrast, figure 6 reports the vote share of Federal Right wing parties, without reporting any marked changes over time. This could be because Batasuna is a leftist party and Federal Leftist parties are slightly more sympathetic to Basque Nationalist positions than Right Wing Federal parties, which tend to seek confrontation. Yet, the magnitude of these changes is quite small.

Figure 7 reports the vote share of other local independent parties, that cannot be classified along the usual ideological axis. The figure shows that in those municipalities in which Batasuna was banned in 2007, the vote for these parties increased, and this effect persisted after the lifting the ban, both in 2011 and in 2015.

¹⁴An alternative simpler explanation would be an exogenous aggregate popularity shock

Figure 1: Turnout

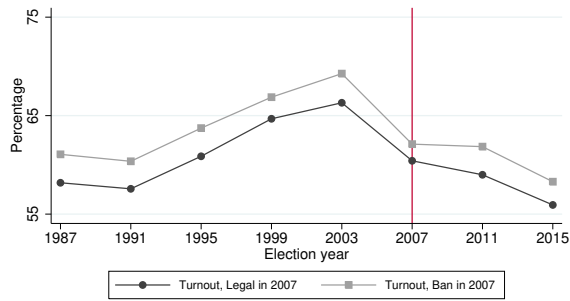


Figure 2: Null votes

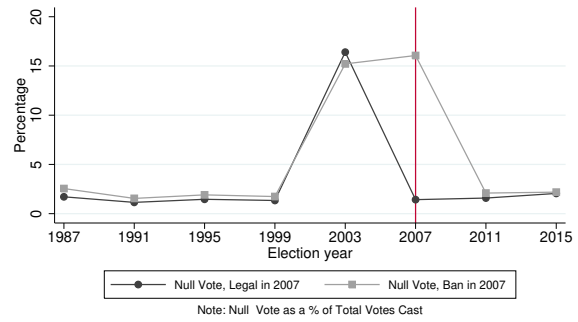


Figure 3: Batasuna votes

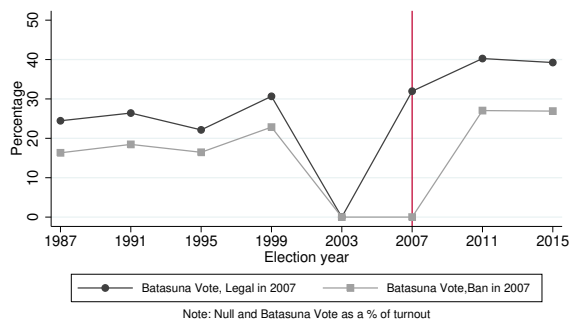


Figure 4: Basque Nationalist votes

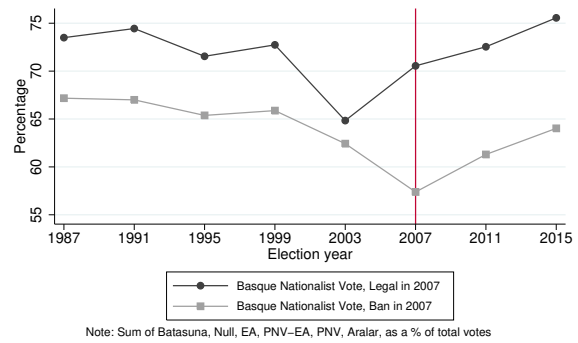


Figure 5: Federal Left

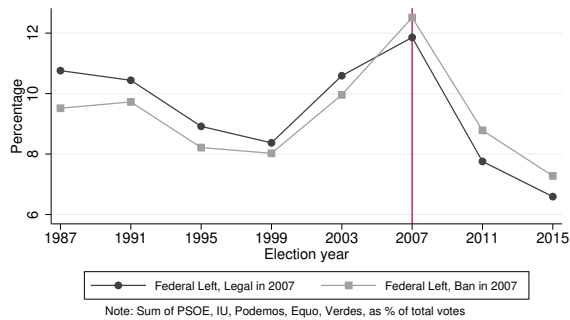


Figure 6: Federal Right

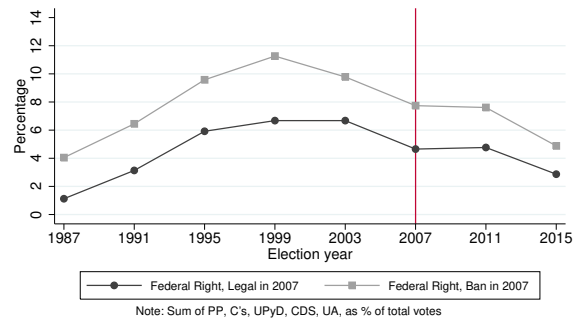
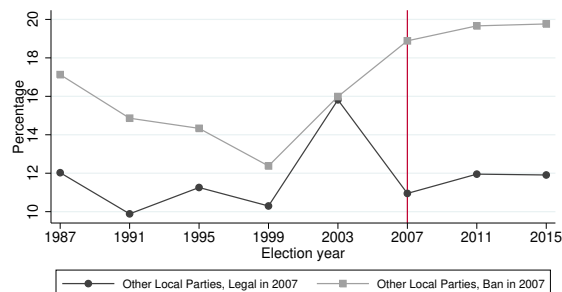


Figure 7: Other local parties



These figures show that the municipalities where the ban lasted for one electoral term (i.e., ban in 2003 but not in 2007) presented similar outcomes to those where the ban lasted for two elections (2003 and 2007). Specifically, they seem to react in a very similar fashion to aggregate popularity shocks (i.e., they seem to be following parallel trends). Because I observe support for Batasuna before and after the ban in all municipalities, and the length of the ban varies across municipalities, conditional on a parallel trends assumption a difference-in-differences estimate recovers a causal effect of the differential length of the ban on the electoral success of Batasuna. In order to investigate the relationship between the 2007 ban on Batasuna and subsequent electoral outcomes, I therefore estimate the following equation:

$$Outcome_{mt} = \alpha_m + \delta_t + \beta \left(I(\text{Post 2007})_t \times I(\text{Ban in 2007})_m \right) + \epsilon_{mt}$$

where $Outcome_{mt}$ is an electoral outcome in municipality m in election t , α_m and δ_t are municipality and election fixed effects, and $I(\text{Post 2007})_t \times I(\text{Ban in 2007})_m$ is a dummy variable equal to one for the treated municipalities after the 2007 ban. The main outcome of interest is the level of electoral support for Batasuna. In all regressions, 2007 is excluded from the sample, because Batasuna had a different legal status across treated and control municipalities, and, hence, differences in support for Batasuna cannot be directly compared. In a heterogeneous treatment effects framework, under the assumption that treated and control municipalities would have followed parallel trends in the absence of the 2007 ban, difference-in-differences estimates an average treatment effect on the treated (ATT). This is an interesting treatment effect, because it is the treatment effect for a subset of municipalities with a considerable level of support for Batasuna. Because Batasuna was banned in all municipalities in 2003, this is an intensive margin rather than an extensive margin effect.

It should be stressed that the identification assumption does not require the 2007 ban to be randomly assigned across municipalities, but to depend on time-invariant unobservables. The 2007 ban was heterogeneous because the public prosecutor failed to find enough information before the election to demonstrate that the electoral lists presented by EAE-ANV were actually Batasuna lists. After the electoral lists were presented, they were banned wherever the connection with Batasuna was most evident (i.e., those including high numbers of former

Batasuna candidates or councilors). Figure 3 shows that municipalities that escaped the 2007 ban had a higher level of pre-ban support, suggesting that this allowed Batasuna to find *clean* candidates more easily, which ended up increasing their probability of escaping the ban. Ex-ante, it was not clear how including Batasuna candidates in the list would play a role, because the precise criterion was established ex-post and because this was the first time that an organization had been banned heterogeneously across municipalities. However, given the context of continuous bans, appearing to be as different as possible from Batasuna on paper was a very likely aim when drawing up the electoral lists. Yet, this seems orthogonal to pre-treatment trends. A visual inspection of figures 1 to 7 shows that treated and control municipalities had followed parallel trends for a long period of time, indicating that they tended to react similarly to aggregate shocks. Hence, if the ability to find *clean* candidates was related to time invariant unobservables, these are absorbed by municipality fixed effects, and difference-in-differences estimates the causal effect of interest.

However, it is interesting to consider why this assumption might be violated. One possibility is that the aim of the Courts was to ban municipalities where support for Batasuna was growing. In this case, difference-in-differences estimates would be biased towards finding a positive effect of the ban on support for Batasuna. Another possibility is that Batasuna was better able to find clean candidates in those places where its support was increasing (i.e., not only where it was higher). In this case, the estimates would be biased towards finding a negative effect of the ban on support. Although it is not possible to observe electoral support immediately before Batasuna presented its lists, in the section on identification threats I test for the existence of differential trends in support for Batasuna across treated and control municipalities in the regional elections that took place in 2005, just two years before the 2007 ban. Moreover, I exploit the criteria employed by the public prosecutor to denounce EAE-ANV lists, based on the number of former Batasuna candidates and councilors, to provide instrumental variables estimates of the effect of the ban. These estimates rely on a weaker identification assumption but have low statistical power, because they combine triple differences with imperfect compliance and a relatively small sample. The instrumental variables identification assumption and the results, which are qualitatively similar to the difference-in-differences estimates, are discussed in Appendix B.

In the estimation of the baseline difference-in-differences, to compare municipalities in which the party had a more similar local infrastructure before the ban, I group municipalities according to whether Batasuna presented lists in each of the local elections before the ban and include year fixed effects for each of these groups, namely pre-ban presence by election fixed effects. The largest group is the one for which Batasuna presented lists in every election between 1987 and 1999.¹⁵ In 2003, I measure support for Batasuna with the share of null votes, and to be consistent across years, the dependent variable -support for Batasuna- is defined as $\frac{Batasuna\ Votes + Null\ Votes}{Total\ Votes}$. Arenas (2016) shows that under the 2003 ban, null votes only increased in municipalities where Batasuna had obtained votes in the past, and that they represented at least two-thirds of the votes that they would have received in a regular election, indicating that this is a good measure of support. Recall that in all regressions, 2007 is excluded from the sample, because support is not comparable because of the different legal status, which is precisely our treatment of interest.

Tables 1 reports the difference-in-differences estimates of the effect of the 2007 ban on the outcomes in figures 1 to 7. Table A1 in the Appendix reports descriptive statistics by treatment group and period. As suggested by the figures, the results in the top panel show that a longer ban had no effect on turnout nor on null votes, but that it had a negative effect of around 6pp on Batasuna’s vote share, which translates into a negative effect of a similar size on the vote share of nationalist parties overall. The magnitude of the effect is important: it represents around 15% of the average post-ban support for Batasuna, or almost 0.3 standard deviations. This compares to 8.7% of all the nationalists’ vote share, or 0.2 standard deviations.

The result on turnout is important for the interpretation of these results. One possible concern about interpreting the results as a negative effect of the ban on support for Batasuna’s positions is the party’s move towards a more moderate stance in relation to ETA after the lifting of the ban. For instance, Gould and Klor (2010) show that in Israel terrorism makes voters more likely to vote for the right, but at the same time it shifts their preferences to the

¹⁵Groups are defined for all possible combinations of presenting electoral lists between 1987 and 1999. For instance, municipalities with no lists until 1991 but with lists in 1995 and 1999 are grouped together, as are municipalities with lists only in 1991, municipalities present in all elections, municipalities present in all elections but 1995, and so on.

left. In this setting, an analogous outcome would be people no longer voting for Batasuna because they find them too moderate. Although without individual level data I cannot completely rule this possibility out, given that Batasuna remains the most extreme platform, the result that a longer ban does not affect turnout, combined with the parties' vote share patterns, suggests that this is unlikely.

The bottom panel of table 1 reports a small positive effect on the vote for the Federal left, a slightly smaller negative effect on the vote for the Federal right, and an increase in votes for local parties, as suggested by the previous figures.

Table 1: Effects of the 2007 ban, local elections

	Turnout (1)	Null Votes (2)	Batasuna Support (3)	All Basque Nationalists (4)
$I(\text{Post 2007}) \times I(\text{Ban in 2007})$	-0.00306 (0.00558)	-0.00130 (0.00461)	-0.0636 (0.0187)	-0.0585 (0.0253)
Municipality FE	✓	✓	✓	✓
Pre-ban presence \times election FE	✓	✓	✓	✓
N	1735	1735	1735	1735
Municipalities	251	251	251	251

Standard errors clustered at the municipality level in parentheses. Support for Batasuna is measured as the sum of Batasuna votes and Null votes, as a fraction of total votes. Nationalist support is the sum of Support for Batasuna, PNV, EA, PNV-EA, and Aralar, as a fraction of total votes. In all regressions, 2007 is excluded from the sample. All regressions control for a cubic in municipality size.

	Federal Left (1)	Federal Right (2)	Other local parties (3)
$I(\text{Post 2007}) \times I(\text{Ban in 2007})$	0.0132 (0.00666)	-0.00865 (0.00414)	0.0572 (0.0284)
Municipality FE	✓	✓	✓
Pre-ban presence \times election FE	✓	✓	✓
N	1735	1735	1735
Municipalities	251	251	251

Standard errors clustered at the municipality level in parentheses. Federal left is the sum of PSOE, IU, Podemos, Los Verdes, and Equo, as a fraction of total votes. Federal Right is the sum of PP, Ciudadanos, Unidad Alavesa, Centro Democrático y Social, and UPyD. Other local parties are votes that are neither for Nationalists, Federal Left or Federal Right. In all regressions, 2007 is excluded from the sample. All regressions control for a cubic in municipality size.

Identification threats

The interpretation of the previous results as causal effects relies on the common trends assumption. This assumption is fundamentally untestable, but pre-treatment differences can nonetheless be informative about its plausibility. To formally test for pre-treatment differences in trends, table A2 in the Appendix reports placebo regressions, that include a lead treatment indicator which is never significant and mostly close to zero, indicating that the estimates in table 1 are not driven by differential pre-ban trends, as suggested by the figures.

In the case of null votes under the 2003 ban, table A8 in the Appendix further tests for differences in null votes between treatment and control municipalities. The estimates in the first column, accounting for municipality and election fixed effects, indicate that in 2003 null votes increased slightly less in treated municipalities, as suggested by figure 2. Note however that municipalities in which Batasuna was banned in 2007 also presented a lower level of support for Batasuna before the ban, and, hence, one would also expect fewer null votes in those municipalities in 2003. In fact, the difference in null votes in 2003 is smaller than the pre-ban difference in support for Batasuna between treated and control municipalities, which means that treated municipalities had lost less support in 2003, compared to municipalities escaping the ban in 2007.

Column 2 reports estimates that include pre-ban presence by year fixed effects, as in the baseline specifications throughout the paper, with the aim of comparing municipalities with more similar pre-ban strengths. After conditioning on these, municipalities with a longer ban received slightly more null votes in 2003, but again this has to be compared with the pre-ban difference in support after conditioning on pre-ban presence by election fixed effects. This is done in column 3, which reports difference-in-differences estimates on support for Batasuna (defined as $\text{Batasuna Votes} + \text{Null Votes}$) for the full sample, including a lead treatment indicator to test for differences under the 2003 ban. The results in column 2 and 3 indicate that the difference between treated and control municipalities under the 2003 ban is small, which is consistent with the common trends assumption.

In another test of pre-treatment trends, column 2 of table A5 in the Appendix examines the effect of the 2007 ban in local elections on support for Batasuna in the Basque regional elections, reporting that in the 2005 regional election, just two years before the 2007 ban, treated and control municipalities were still performing similarly. Columns 1 and 2 in table A5 report that, after 2007, Batasuna lost support (around 1.2pp) in treated municipalities in the regional elections.

Another identification threat would be the existence of compound treatment effects. For instance, the changes in Batasuna’s platform after the end of the ban could have a differential appeal across treatment and control municipalities. As a robustness check, table A10 in the Appendix includes estimates that allow for heterogeneous post-ban effects, depending on a number of covariates that could plausibly induce heterogeneous effects. Across columns, the estimates allow for heterogeneity in post-ban election effects by municipality size, (log) terror attacks in the 10 years before the ban, (log) street violence attacks in the year before the ban; by pre-ban local electoral outcomes, including Batasuna’s vote shares, the sum of Basque nationalists, and the sum of federal parties; and by pairwise interactions between each of these six variables. The results indicate that the estimates are robust to these heterogeneous time effects.¹⁶

Another source of violation of the common trends assumption concerns the definition of the outcome of support for Batasuna. As explained in the section on the ban on Batasuna, after the ban, Batasuna re-emerged as Bildu, a new coalition led by Batasuna, but in conjunction with smaller parties that enjoyed little electoral success before the ban, such as EA (social democrat Basque nationalists), and Alternatiba (a splinter group from IU, federal leftists). This could be problematic if there had also been a differential increase in support for some of Batasuna’s new partners at the same time of the ban. The effect on the Basque nationalist vote share overall, which includes almost all of these partners (and also the large Christian-democrat Basque nationalists (PNV)), suggests that this is not a concern. Nonetheless, I also present entropy balancing matching estimates (Hainmueller, 2012) to further deal with this. Entropy balancing calibrates municipality weights so that

¹⁶Street violence attacks are only available since April 2006. See the section on street violence for further details.

the re-weighted treatment and control group satisfy a set of pre-specified balance conditions. More specifically, I use entropy balance weights so that the pre-treatment support for each political party is on average the same across treatment and control municipalities. With these weights, any differential increase in support for any political party would have the same effect in treated and control municipalities, on average.¹⁷ The entropy balancing estimates in table A9 in the Appendix are very similar to those in table 1, suggesting that this cannot explain the results.

Heterogeneity and mechanisms

A party ban is a repressive measure that prevents a specific political party from contesting elections and obtaining any of the benefits from being in office or holding seats in parliament. Hence, previous studies of the effects of party bans have primarily considered two sets of possible mechanisms. A first set of mechanisms concerns citizens' reactions to the ban as a policy. These are assumed to be negative because of the democratic contradiction that it might entail, possibly generating sympathies towards the party among people who are simply opposed to the ban. Given the overall negative effect of the ban on support for Batasuna, these mechanisms do not seem to be playing a central role in this case. A second set of mechanisms concerns the relative disadvantages faced by the targeted party, given its inability to contest elections and hold parliamentary (city council) representation. This might be associated with a loss of support because of (1) the direct material disadvantages of not holding parliamentary representation and its effect on the party's mobilization capacity, and (2) voters' increased information about competing politicians and parties that may not have been competitive or elected otherwise.

It is difficult to disentangle empirically each of these mechanisms, which have not been

¹⁷More precisely, I balance the level of support for a number of political parties in each of the four pre-treatment elections (2003, 1999, ..., 1987). These are the vote shares of Batasuna; the vote shares of the Basque Nationalist Party (PNV, Basque Christian Democrats), the vote shares of Eusko Alkartasuna (EA, Basque Social-Democrats, which becomes part of Bildu); the vote shares of EA-PNV coalitions, that take place in some municipalities in some pre-treatment elections, the vote shares of IU (Federal leftist, a splinter of which becomes part of Bildu); the vote shares of Aralar, a small leftist independentist moderate party that joined Bildu (which pre-treatment only contested the 2003 local election); the vote shares of the PP (the Federal conservatives), and the vote shares of the PSOE (the Federal social democrats); and also dummies for Batasuna being present in each of the pre-ban elections.

formalized in previous qualitative party ban studies. A particular feature of the ban on Batasuna is that when it did not apply to all municipalities, the party continued to contest elections in those places where it was banned by calling for a null vote. Hence, it is possible to observe the immediate loss in support for the party under the ban, because null votes are otherwise typically very close to zero. This is interesting because an understanding of the timing of the effect of the ban can be informative about the mechanisms. Here, given the sudden nature of the 2007 ban, any difference in the party’s mobilization capacity and the material benefits of holding city council seats should emerge gradually over time. On the other hand, if the main mechanism involves voters learning about new parties, one would expect a significant correlation between the very immediate effect of the ban on support for Batasuna in 2007 (when Batasuna called for a null vote in treated municipalities) and the effect once the ban was lifted. There are two explanations for this: first, that when more Batasuna voters abandon the party under the ban, the composition of the remaining parties on the city council is less likely to be the same: recall, those gaining the most from the ban are small local parties, with less known platforms, as shown in figure 7 and table 1. And second, that because under the ban voting for Batasuna does not translate into seats, instrumental voters have incentives to learn about new parties, which could result in an immediate vote loss that persists on the lifting of the ban. These mechanisms would lead to a positive correlation between the immediate and the long-run losses in support resulting from the ban.

Figure 8 maps the immediate loss in support in 2007, and the loss in support after the lifting of the ban (excluding 2007) across municipalities. The pattern is similar, suggesting that those municipalities where voters immediately switched allegiance to another party were also those municipalities that lost support on the lifting of the ban. Column 1 in table 2 reports estimates of the effect of the ban that control for an interaction of the average pre-ban vote share of Batasuna and a post-2007 dummy. This does not affect the results, indicating that the average effect of the ban is orthogonal to the average pre-existing strength of the party. Column 2 in table 2 further controls for an interaction of the 2007 vote share (driven by votes for Batasuna in those municipalities where it was legal, and by null votes in those where it was banned) and a post-2007 dummy. Note that this is a bad control (i.e., given

that it is in itself an outcome of the ban), and, hence, our interest in this specification is descriptive rather than causal. The results indicate that those municipalities where Batasuna immediately lost support were also those municipalities where Batasuna lost support on the lifting of the ban. This result is interesting because the immediate loss in support occurred before any difference in the direct disadvantages from not being on the city council could make themselves manifest as a result of the 2007 ban, which is inconsistent with this factor being very important for the estimated effect.¹⁸ Rather, this pattern is more consistent with voter learning about new candidates.

I also examine in which municipalities Batasuna was more or less likely to lose votes, studying heterogeneous effects of the ban across municipalities depending on their level of pre-ban violence and on voters' incentives to abandon Batasuna under the ban. The literature on counter-terrorism has emphasized that a backlash is more likely with measures that inflict damage to a broader population than that of the terrorist group, which might perceive the measure as illegitimate (Bueno de Mesquita and Dickson, 2007; Benmelech et al., 2015). In this case, one might expect voters in municipalities that have not witnessed many episodes of terror to consider the ban as being less legitimate, and, hence, to maintain their vote for Batasuna. A second source of heterogeneity concerns voter incentives when votes for Batasuna are counted solely as null votes. In this case, one might expect voters to stick more firmly to their vote when such a decision is less consequential, for instance when it does not make it easier for antagonist parties to gain power. Table A3 in the Appendix reports descriptive statistics showing that in those municipalities in which Batasuna was banned in 2007, the number of terror attacks in the previous decade had been 13% lower than in the control group, that the margin between nationalist and federal parties in 2003 was a little narrower, and that the population size was smaller.

Table 3 reports estimates that investigate these questions by including a number of interaction terms. Column 1 reports estimates on the immediate vote loss (i.e., support for Batasuna under the 2007 ban), while column 2 provides estimates on the vote loss once

¹⁸An effect through candidates' incentives to join Batasuna also seems less consistent with the data because one would also expect this to emerge progressively over time. In addition, the ban also has an effect on support for Batasuna in regional elections, where treated and control municipalities vote for the same candidatures, as shown in table A5, and also in regional elections the immediate vote loss predicts the effect of the ban.

Figure 8: 2007 ban, immediate and long run loss in support

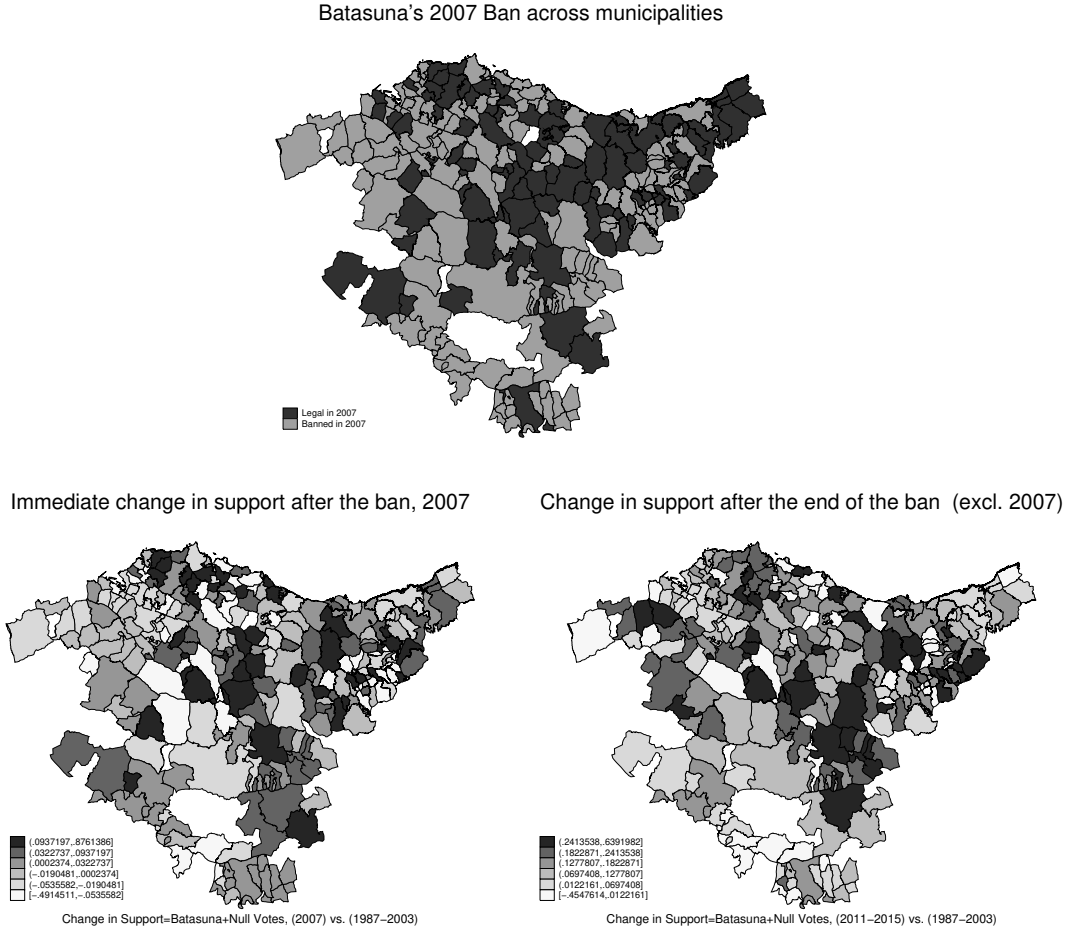


Table 2: Accounting for voting behavior under the ban

	Batasuna Support	
	(1)	(2)
$I(\text{Post } 2007) \times I(\text{Ban in } 2007)$	-0.0674 (0.0185)	-0.00162 (0.0210)
Municipality FE	✓	✓
Pre-ban presence \times election FE	✓	✓
$I(\text{Post } 2007) \times \text{Pre-ban Support}$	✓	✓
$I(\text{Post } 2007) \times \text{Support in } 2007$		✓
N	1733	1733
Municipalities	250	250

Standard errors clustered at the municipality level in parentheses. Support for Batasuna is measured as the sum of Batasuna votes and Null votes. Pre-ban support is the 1987-2003 average. In all regressions, 2007 is excluded from the sample. All regressions control for a cubic in municipality size.

Table 3: Correlates of vote loss

	Batasuna Support	
	(1)	(2)
$I(\text{Year} \geq 2007) \times I(\text{Ban in 2007})$	-0.138 (0.0209)	-0.0731 (0.0214)
$I(\text{Year} \geq 2007) \times I(\text{Ban in 2007}) \times \ln(1 + \# \text{Error Attacks})$	-0.0451 (0.0211)	-0.0686 (0.0337)
$I(\text{Year} \geq 2007) \times I(\text{Ban in 2007}) \times \text{Close}$	-0.104 (0.0496)	-0.0831 (0.0462)
$I(\text{Year} \geq 2007) \times I(\text{Ban in 2007}) \times \ln(\text{Population in 2007})$	0.0515 (0.0155)	0.0395 (0.0166)
Sample	Until 2007, included	Full (excluding 2007)
Municipality FE	✓	✓
Pre-ban presence \times election FE	✓	✓
$I(\text{Year} \geq 2007) \times \ln(1 + \# \text{Error Attacks})$	✓	✓
$I(\text{Year} \geq 2007) \times \text{Close}$	✓	✓
$I(\text{Year} \geq 2007) \times \ln(\text{Population in 2007})$	✓	✓
$I(\text{Year} \geq 2007) \times \text{Support in 2003}$	✓	✓
$I(\text{Year} \geq 2007) \times I(\text{Ban in 2007}) \times \text{Support in 2003}$	✓	✓
N	1477	1725
Municipalities	249	249

Standard errors clustered at the municipality level in parentheses. All interacted variables are demeaned Support for Batasuna is measured as the sum of Batasuna votes and Null votes. $\#$ Error attacks is the number of ETA attacks in the municipality in the 10 years prior to the 2007 ban. Close is a dummy equal to one if in 2003, the absolute difference in vote share between Federal and Basque parties was below 10%. Support in 2003 refers to support for Batasuna. All regressions control for a cubic in municipality size.

the ban had been lifted (i.e., excluding 2007). The interacted variables are demeaned so that the main term of the effect of the ban can be interpreted at their mean values. Note that the regression also controls for interactions with a post-treatment dummy. First, the results indicate that in those municipalities where ETA had been more active in the decade before the 2007 ban, the loss in support was more pronounced, which is consistent with the idea that more indiscriminate policies are perceived as less legitimate.¹⁹ Second, the results indicate that wherever the margin in the most recent local election between Basque Nationalist and Federal parties had been close (i.e., less than 10pp), the vote loss was larger, which is consistent with the idea that pivotality considerations might be relevant. It is interesting to note that the immediate and the long-run heterogeneous effects of the ban are similar. This is consistent with the previous finding that the immediate vote loss due to the

¹⁹The source of information about ETA attacks is the Global Terrorism Database.

2007 ban is a good predictor of the vote loss on the lifting of the ban. Again, this pattern is more consistent with voters learning about competing parties, which could be proportional to the immediate loss (either because of voter learning before the 2007 election, or during the electoral term), and less consistent with the direct benefits from holding seats in the city council, which one would expect to emerge progressively over time. Hence, this suggests that party bans may be used by incumbents to reduce their challengers' strength by triggering an electoral reshuffling, with voters learning about new parties and potentially switching.

Street violence

In addition to ETA's terrorist attacks, which killed more than 800 people, its supporters perpetrated thousands of low-intensity violent attacks while the group remained active. These attacks, known as *Kale Borroka* (or street fights), involved primarily the burning of public buses, private vehicles and trash containers, and incendiary bomb (Molotov cocktail) attacks on court buildings, banks, political party headquarters and police stations, typically perpetrated at night or following demonstrations. de la Calle (2007) studied their determinants between 1996 and 2000, and describes it as a persistent type of non-organized violence (that is, no permanent organizations claimed responsibility for the attacks), which tended to occur in the more polarized municipalities. All political parties except Batasuna condemned the violent squads involved in the *Kale Borroka*, forming part of which was typically the first step towards becoming an ETA militant. In 2000, Spanish-based insurance companies paid out almost 9 million euros because of the acts of violence of the previous three years (de la Calle, 2007). In 2002, the courts held Batasuna civilly responsible for a number of these violent street attacks, and condemned them to pay part of their costs, arguing that the finances of Batasuna and the *Kale Borroka* were jointly determined.²⁰ Given that under Spanish law, such attacks are considered acts of terrorism, in what follows, I use the terms street violence and street terrorism interchangeably. Between 2006 and 2008, more than 350 acts of street terrorism took place in the Basque Country, and in earlier periods they had been even more frequent.

²⁰El País, July 2002.

The effect of the party ban on this type of violence is interesting for two reasons. First, after accounting for municipality and time fixed effects, it is a measure of the intensive margin of support for the targeted movement by violent means. Second, it is informative of the reaction of a terrorist organization to a measure of political repression. Batasuna and its supporters combined both political and violent mobilizations to maximize the threat of revolution, subject to a certain budget constraint.²¹ A party ban might be seen as a tax on political mobilization, and given the central role played by Batasuna in the movement, one would expect it to have large income effects - reducing overall the strength of the movement. These income effects would go in the opposite direction to that of the substitution effect arising from the relative decrease in the price of violent mobilization with respect to the price of political mobilization due to the ban. However, if the ban were to make the aggrieved population more willing to support the targeted movement, increasing their ideological motivation, this could also lead to an income effect of opposite sign, a *motivation effect*. Bueno de Mesquita (2005) fully models the strategic interaction between a terrorist organization, a government, and a population of potential terrorist volunteers, where counterterrorism reduces the ability of terrorists to carry out attacks, by weakening their infrastructure, while at the same time it diminishes their economic opportunities, by reducing their opportunity costs, albeit while increasing their ideological motivation.²²

I estimate the effect of the May 2007 ban on street violence using detailed data at the municipality level for street terrorism episodes perpetrated between April 2006 and June 2008, as reported in Buesa (2008), based on media reports. Table A6 in the Appendix reports descriptive statistics by treatment group and period.²³ Figure 9 displays the average monthly number of street terrorism episodes, by treatment group. Figure 10 shows the same data but averaged over quarters rather than months. The figures show that before the ban was extended in May 2007 the trends in street terrorism ran parallel, and that the extension of

²¹For instance, Aidt and Franck (2015) show that low intensity violence has the potential to trigger major political changes.

²²This framework is useful for cases like the Basque Country, where street violence was mostly one-sided. However, in other cases, for instance in Colombia, Government paramilitaries could play a role, using violence as a de facto reaction to changes in the de jure power of traditional outsiders (Fergusson et al., 2017).

²³I include all the episodes in Buesa (2008) involving petrol bombs, damage and destruction with blunt objects, or heavier attacks. Demonstrations, or attacks with eggs or paints are excluded from the sample. The data in de la Calle (2007) cover only the 1996-2000 period and are based on a collection of media reports.

the ban triggered a peak in street terrorism in the very short run. However, after this spike, the ban led to a slight decline in violence in treated municipalities, especially compared to pre-ban differences, the overall effect being ambiguous.²⁴

Figure 9: Effect of the 2007 ban on street violence, monthly

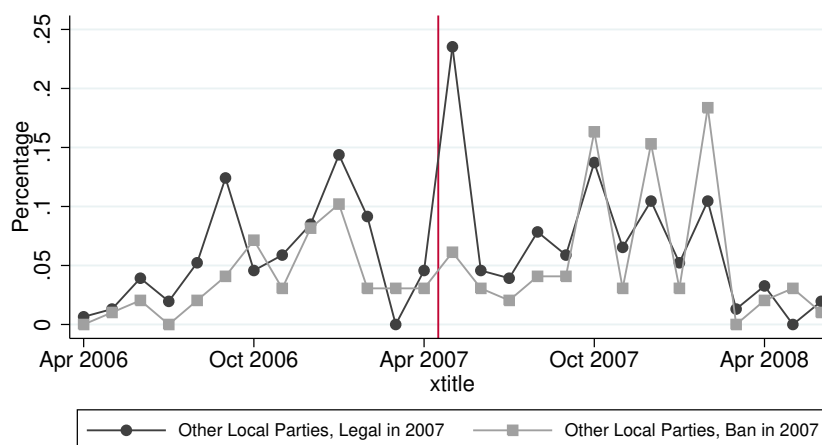
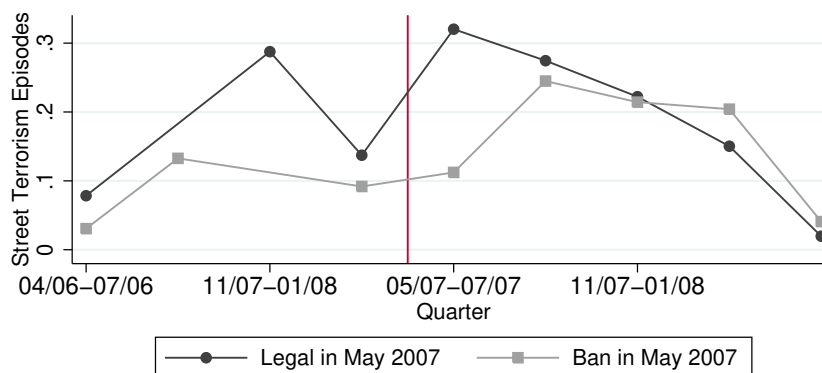


Figure 10: Effect of the 2007 ban on street violence, quarterly



²⁴The formal denunciation was made by the public prosecutor on the 3rd May 2007, and the Supreme Court took its decision on the 6th May. The first episode of violence in May 2007 was recorded as taking place on the 7th May.

Table 4 reports difference-in-differences estimates. In the top panel, the dependent variable is the monthly number of street terrorism episodes in a municipality; in the bottom panel, a dummy for whether there was a street terrorist episode in a given municipality in a given month. All regressions include municipality fixed effects and period (month \times year) fixed effects by pre-ban presence of Batasuna. The regression results in column 1 of both panels show that wherever the ban was extended, the probability of an attack in the following month increased by almost 10pp, with 0.2 more attacks being carried out, on average, in those municipalities. With around 150 treated municipalities, this implies a spike of around 30 attacks in the month after the ban. However, if I consider the full post May 2007 period, the overall effect is close to zero. The results in column 3, which exclude May 2007 from the sample, show that this is because after the very short run spike, the extension of the ban slightly decreased the monthly probability of a street terrorist attack in treated municipalities in the remaining months by 1.3pp, although this is not precisely estimated.

Table 4: Effects of the 2007 ban on street violence

	ST Episodes		
	(1)	(2)	(3)
$I(\text{Post May 2007}) \times I(\text{Ban in 2007})$	0.200 (0.0819)	-0.00430 (0.0177)	-0.0200 (0.0166)
Sample	Until May 2007	Full	Except May 2007
Municipality FE	✓	✓	✓
Pre-ban presence \times Month FE	✓	✓	✓
N	3514	6777	6526
Municipalities	251	251	251

Standard errors clustered at the municipality level in parentheses. The dependent variable is given by the monthly number of street terrorism attacks.

	I(ST Episode)		
	(1)	(2)	(3)
$I(\text{Post May 2007}) \times I(\text{Ban in 2007})$	0.106 (0.0371)	-0.00475 (0.00971)	-0.0133 (0.00988)
Sample	Until May 2007	Full	Except May 2007
Municipality FE	✓	✓	✓
Pre-ban presence \times Month FE	✓	✓	✓
N	3514	6777	6526
Municipalities	251	251	251

Standard errors clustered at the municipality level in parentheses. The dependent variable is a dummy equal to one if there is one or more episode of street terrorism in a month.

The results are consistent with a positive motivation effect of the ban that vanishes quickly, in line with the findings of Jaeger et al. (2012) of a fleeting radicalization of the Palestinian population after Palestinian fatalities. Table A7 in the Appendix reports the results of a placebo test, indicating that shortly before May 2007 treated and control municipalities were performing similarly, as suggested by figures 9 and 10.

General equilibrium effects and external validity

One limitation of this case study is the possible existence of general equilibrium effects. The first of these, common to most public policy evaluations, concerns the difference between the effect of the implementation of the ban and the effect of a credible announcement of a reaction function by the government that could eventually lead to a ban (i.e., the equilibrium effect). For instance, Dragu (2011) and Dragu (2016) show how reducing privacy or constraining free speech in response to terrorist attacks could increase terrorism in the first place, by changing the security agencies' stakes for terrorism prevention or by possibly increasing the terrorist organization's initial stake for a successful terrorist attack. Hence, one avenue for future research is identifying an empirical setting with potential for capturing how the existence of the ban might affect, in the future, the probability of observing a ban, or terrorism, in the first place, through its effects on terrorism or political behavior.

A second type of general equilibrium effect, more specific to this setting, concerns the existence of externalities across municipalities, which could violate the Stable Unit Treatment Value Assumption (SUTVA) (Rubin, 2005) that is required for a causal interpretation of the estimates in this paper. In the ideal experiment, party bans would be randomly assigned across independent party systems. However, this is not feasible, but without random assignment, it is difficult to find plausible control groups in independent party systems. This paper overcomes this problem by exploiting the differential intensity of a given ban within a party system, across municipalities that are comparable and which provide a plausible control group. The cost of this approximation is the possibility of spillovers between municipalities. For instance, the effect on party mobilization capacity and resources that would lead the ban to reduce political support could be attenuated by the presence of Batasuna in some municipalities

in 2007, which would bias the results towards finding no effect of the ban, because the party could be mobilizing and spending resources from control to treated municipalities. Nevertheless, given the large estimate of the effect of the 2007 ban, this is unlikely to be a concern. On the other hand, the population in municipalities where Batasuna was legal in 2007 could still react negatively to the ban as a policy because Batasuna was banned in some other municipalities, which could increase support for Batasuna in control municipalities. This would bias the results instead towards finding effects of the ban that are overly negative. While this cannot be fully ruled out, the immediate effect of the 2007 ban clearly seems driven by a decrease in support in treated municipalities (i.e., by null votes in treated municipalities being substantially lower than Batasuna votes in pre-ban elections) rather than by an increase in support for Batasuna in control municipalities. This immediate drop is the main predictor of the long-run effect, which suggests that this explanation is not very likely.

Finally, the focus of this paper on the Basque conflict raises concerns regarding the external validity of the results. A first external validity concern is that the ban of Batasuna was a temporary ban. Nonetheless, it turns out that this is no exception, and often the targeted political movement manages to return, following periods of variable length, under a similar guise.²⁵ For instance, the Sinn Féin was banned between 1956 and 1974 in Northern Ireland, and the Turkish Islamist Welfare Party was succeeded by the Islamist Virtue Party shortly after its ban in 1998. In Israel, the far right Kach Kahanist party, which was banned in 1994, was succeeded by Otzma Yehudit in 2012, and tolerated by the courts. The Communist Party of Germany (KPD) was banned in 1956 and was reconstituted in 1968 as the German Communist Party (PKD).

A second external validity concern is that the 2007 ban on Batasuna is informative of the intensive margin effect of an additional or longer ban rather than of the extensive margin, and Batasuna is a party with a fairly solid support base. Table A11 in the Appendix reports heterogeneous effects of the ban by the average level and variance of pre-ban support. Note that the interacted variables are demeaned so that the coefficient of the main term can be interpreted as the effect at the mean. While the heterogeneous effect by average pre-ban support is imprecisely estimated and the effect at the mean is very close to the baseline

²⁵See Bourne (2018) for a comprehensive list of parties subject to ban proceedings in Europe (1945–2015).

estimate, the heterogeneous effect by the variance of the pre-ban support suggests that the loyalty of the support base, or the voters' eagerness to switch parties, is significantly related to the effect of the ban.²⁶ However, this case study cannot identify possible non-linearities concerning ban length and electoral support.

Given the inevitable trade-off between identification via case studies and external validity, future research focused on other conflicts should be able to shed more light on this question and its associated mechanisms, because empirical studies of conflict inherently take place in unstable and changing institutional environments, making the results more contingent on a particular context.²⁷ Nevertheless, it should be borne in mind that, overall, the estimates in this paper are consistent with the qualitative study by Bale (2007), who following an analysis of the bans on Batasuna, the Islamist Welfare Party in Turkey, and the far-right Vlaams Blok in Belgium, concludes that the pragmatic predictions that party bans routinely occasion –namely *perversity, futility and jeopardy*, as in Hirschman (1991)– are not borne out in practice.

Conclusions

Party bans are common in many countries, but disentangling their effects from the political circumstances leading to the proscription is not easy. In this paper, I have attempted to circumvent this problem by empirically evaluating the electoral effects of the ban on Batasuna, the political wing of ETA, exploiting the finite and heterogeneous nature of the ban across Basque municipalities. The main result of the paper is that a longer ban had a negative effect on support for Batasuna, the effect being large and persistent for at least two elections following the lifting of the ban. Furthermore, using data on episodes of street violence perpetrated by Batasuna supporters, the results indicate that the extension of the ban for an additional electoral term led to a one month spike in violence in treated municipalities, but it had no effect during the following year. These results are interesting because while party bans

²⁶Nonetheless, for any value of the average pre-ban support and its variance in the sample, the predicted effect of the ban is negative.

²⁷For instance, in this case, the context was a general non-conciliatory stance by the Spanish government and judiciary. Hence, a limitation of this paper is that it is difficult to know whether the effect of a ban would change with a different mixture of repressive and non-repressive government policies.

deprive the extremist movement of organizational resources and social visibility, they might encounter strong principled opposition in public opinion because of the contradiction with democratic principles. The interpretation of the results as a causal effect of the ban relies on the assumption that both groups of municipalities would have followed parallel trends in electoral outcomes in the absence of the heterogeneous length of the ban. An inspection of pre-treatment trends suggests that this is a plausible assumption.

Although it is difficult to disentangle the specific mechanisms that led to the loss of support for the party, a number of heterogeneity analyses point towards the importance of voter learning about parties that otherwise would have not been competitive or elected. This is because the immediate effect of the ban on support for Batasuna, which is measurable because Batasuna contested the elections under the ban by calling for a null vote, is a very good predictor of the loss in support following the lifting of the ban. A greater immediate loss in support is associated with a greater change in the make-up of the city council, and instrumental voters have an incentive to learn about new parties under the ban, which would correlate both with the immediate and with the long-term loss in electoral support. If instead the effect of the ban were due to the material advantages of being on the city council, one would not expect the effect to emerge immediately but rather more progressively. The loss in support is also more marked in municipalities where voters are more prone to switching parties, which is also consistent with this interpretation. Hence, overall the results suggest that party bans may be used by incumbents to reduce their challengers' strength by triggering an electoral reshuffling, with voters learning from the new scenario and potentially switching parties. Further examination of these mechanisms is left for future research.

References

- Abadie, A. and J. Gardeazabal (2003). The Economic Costs of Conflict: A Case Study of the Basque Country. *American Economic Review* 93(1), 113–132.
- Aidt, T. S. and R. Franck (2015). Democratization under the threat of revolution: Evidence from the Great Reform Act of 1832. *Econometrica* 83(2), 505–547.
- Ansolabehere, S. and M. S. Puy (2016). Identity voting. *Public Choice* 169(1), 77–95.

- Arenas, A. (2016). Sticky votes. *Journal of Economic Behavior & Organization* 132(A), 12 – 25.
- Balcells, L. (2012). The consequences of victimization on political identities: Evidence from Spain. *Politics & Society* 40(3), 311–347.
- Bale, T. (2007). Are bans on political parties bound to turn out badly? *Comparative European Politics* 5(2), 141–157.
- Barros, C. P., G. M. Caporale, and L. A. Gil-Alaña (2009). Basque terrorism: police action, political measures and the influence of violence on the stock market in the Basque Country. *Defence and Peace Economics* 20(4), 287–301.
- Basedau, M. and A. Moroff (2011). Parties in chains: Do ethnic party bans in Africa promote peace? *Party Politics* 17(2), 205–222.
- Benmelech, E., C. Berrebi, and E. F. Klor (2015). Counter-Suicide-Terrorism: Evidence from House Demolitions. *Journal of Politics* 77(1), 27–43.
- Bourne, A. K. (2018). *Democratic Dilemmas: Why democracies ban political parties*. Routledge.
- Bueno de Mesquita, E. (2005). The quality of terror. *American Journal of Political Science* 49(3), 515–530.
- Bueno de Mesquita, E. and E. S. Dickson (2007). The propaganda of the deed: Terrorism, counterterrorism, and mobilization. *American Journal of Political Science* 51(2), 364–381.
- Buesa, M. (2008). *Actividades terroristas de ETA y política antiterrorista: información estadística*. Cátedra de Economía del Terrorismo, UCM, Link.
- Capoccia, G. (2005). *Defending democracy: Reactions to extremism in interwar Europe*. JHU Press.
- Capoccia, G. (2013). Militant democracy: The institutional bases of democratic self-preservation. *Annual Review of Law and Social Science* 9, 207–226.
- Casal Bértoa, F. and A. Bourne (2016). Prescribing democracy? Party proscription and party system stability in Germany, Spain and Turkey. *European Journal of Political Research*.
- Council of Europe - Venice Commission (1999). *Guidelines on the Prohibition and Dissolution of Political Parties and Analogous Measures*. Strasbourg: Council of Europe.

- Da Silveira, B. S. and J. M. De Mello (2011). Campaign advertising and election outcomes: Quasi-natural experiment evidence from gubernatorial elections in Brazil. *The Review of Economic Studies* 78(2), 590–612.
- de la Calle, L. (2007). Fighting for Local Control: Street Violence in the Basque Country. *International Studies Quarterly* 51(2), 431–455.
- de la Calle, L. and I. Sánchez-Cuenca (2013). Killing and Voting in the Basque Country: An Exploration of the Electoral Link Between ETA and its Political Branch. *Terrorism and Political Violence* 25(1), 94–112.
- Dinas, E., P. Riera, and N. Roussias (2015, 5). Staying in the first league: Parliamentary representation and the electoral success of small parties. *Political Science Research and Methods* 3, 187–204.
- Downs, W. M. (2012). *Political extremism in democracies: Combating intolerance*. Palgrave Macmillan.
- Dragu, T. (2011). Is There a Trade-off between Security and Liberty? Executive Bias, Privacy Protections, and Terrorism Prevention. *American Political Science Review* 105, 64–78.
- Dragu, T. (2016). The moral hazard of terrorism prevention. *Journal of Politics* 1(79).
- Dragu, T. (2017). On repression and its effectiveness. *Journal of Theoretical Politics* 29(4), 599–622.
- Dragu, T. and M. Polborn (2014). The rule of law in the fight against terrorism. *American Journal of Political Science* 58(2), 511–525.
- Fergusson, L., P. Querubin, N. A. Ruiz-Guarin, and J. F. Vargas (2017). The real winner’s curse. *WP, Documentos CEDE 2017-05, Link*.
- Fox, G. H. and G. Nolte (1995). Intolerant democracies. *Harv. Int’l. LJ* 36, 1.
- Gandhi, J. and E. Lust-Okar (2009). Elections under authoritarianism. *Annual Review of Political Science* 12, 403–422.
- Gardeazabal, J. (2011). Linguistic polarization and conflict in the Basque Country. *Public Choice* 149(3), 405–425.
- Garmendia Madariaga, A. (2011). La disolución de formaciones políticas en España: el caso Sortu. *Cuestiones constitucionales* (25), 317–331.

- Gil-Alaña, L. A. and C. P. Barros (2010). A Note on the Effectiveness of National Anti-Terrorist Policies: Evidence from ETA. *Conflict Management and Peace Science* 27(1), 28–46.
- Gould, E. D. and E. F. Klor (2010). Does terrorism work? *Quarterly Journal of Economics* 125(4), 1459–1510.
- Hainmueller, J. (2012). Entropy balancing for causal effects: A multivariate reweighting method to produce balanced samples in observational studies. *Political Analysis* 20(1), 25–46.
- Hirschman, A. (1991). *The Rhetoric of Reaction*. Harvard University Press.
- Horz, C. M. (2016). Party bans. *Mimeo, NYU, Link*.
- Issacharoff, S. (2007). Fragile democracies. *Harvard Law Review* 120, 1405.
- Jaeger, D. A., E. F. Klor, S. H. Miaari, and M. D. Paserman (2012). The struggle for Palestinian hearts and minds: Violence and public opinion in the Second Intifada. *Journal of Public Economics* 96(3), 354–368.
- Larreguy, H. A., J. Marshall, and J. M. Snyder Jr (2018). Leveling the playing field: How campaign advertising can help non-dominant parties. *Journal of the European Economic Association* 16(6), 1812–1849.
- Lupu, N. and L. Peisakhin (2017). The legacy of political violence across generations. *American Journal of Political Science* 61(4), 836–851.
- Minkenberg, M. (2006). Repression and reaction: militant democracy and the radical right in Germany and France. *Patterns of Prejudice* 40(1), 25–44.
- Rozenas, A., S. Schutte, and Y. Zhukov (2017). The political legacy of violence: The long-term impact of Stalin’s repression in Ukraine. *The Journal of Politics* 79(4), 1147–1161.
- Rubin, D. B. (2005). Causal inference using potential outcomes. *Journal of the American Statistical Association* 100(469), 322–331.
- Spenkuch, J. L. and D. Toniatti (2018). Political advertising and election results. *The Quarterly Journal of Economics* 133(4), 1981–2036.
- Van Spanje, J. and C. De Vreese (2015). The good, the bad and the voter: The impact of hate speech prosecution of a politician on electoral support for his party. *Party Politics* 21(1), 115–130.

Party bans: deterrence or backlash? Evidence from the Basque Country

Online Supporting Information

Appendix A

Descriptive statistics and placebo tests, local elections

Table A1: Descriptive statistics, local elections

Variable Name	Treated		Control		Total	
	Pre	Post	Pre	Post	Pre	Post
Turnout	0.643 (0.109)	0.601 (0.109)	0.616 (0.100)	0.575 (0.0928)	0.632 (0.107)	0.591 (0.104)
Batasuna votes	0.148 (0.167)	0.270 (0.211)	0.207 (0.190)	0.397 (0.191)	0.171 (0.179)	0.320 (0.213)
Null votes	0.0463 (0.0898)	0.0214 (0.0191)	0.0444 (0.0904)	0.0182 (0.0150)	0.0455 (0.0900)	0.0201 (0.0177)
Batasuna Support	0.194 (0.165)	0.291 (0.209)	0.251 (0.179)	0.416 (0.190)	0.217 (0.173)	0.340 (0.210)
Vote Share Basque Nationalists	0.656 (0.276)	0.627 (0.292)	0.714 (0.247)	0.741 (0.229)	0.678 (0.266)	0.671 (0.275)
V. Share Federal Left	0.0909 (0.127)	0.0803 (0.108)	0.0981 (0.119)	0.0717 (0.0943)	0.0937 (0.124)	0.0770 (0.103)
V. Share Federal Right	0.0826 (0.123)	0.0625 (0.0958)	0.0474 (0.0800)	0.0382 (0.0628)	0.0688 (0.109)	0.0530 (0.0852)
V. Share Other Local Parties	99.15 (0.279)	99.20 (0.289)	99.12 (0.223)	99.12 (0.203)	99.14 (0.259)	99.17 (0.261)
Population size	10087.3 (38455.7)	10092.2 (37998.0)	6338.5 (13143.4)	6567.9 (12950.6)	8613.9 (31117.6)	8713.5 (30760.9)
Municipalities	152	153	98	98	250	251
N	749	305	485	196	1234	501

Treated: municipalities in which Batasuna was banned in 2007. Batasuna support is the sum of Batasuna votes and Null votes. 2007 excluded from the sample. Standard deviations in parentheses.

Table A2: Effects of the 2007 ban, placebo test, local elections

	Turnout	Null Votes	Batasuna Support	All Basque Nationalists
	(1)	(2)	(3)	(4)
$I(\text{Post } 2007) \times I(\text{Ban in } 2007)$	0.00129 (0.00525)	-0.00581 (0.0101)	-0.0739 (0.0179)	-0.0665 (0.0261)
$I(\text{Post } 1995) \times I(\text{Ban in } 2007)$	-0.00735 (0.00603)	0.00762 (0.00967)	0.0174 (0.0107)	0.0135 (0.0200)
Municipality FE	✓	✓	✓	✓
Pre-ban presence \times election FE	✓	✓	✓	✓
N	1735	1735	1735	1735
Municipalities	251	251	251	251

Standard errors clustered at the municipality level in parentheses. The placebo treatment is equal to zero in 1987, 1991, 1995, and equal to the treatment from 1999 onwards. Support for Batasuna is measured as the sum of Batasuna votes and Null votes, as a fraction of total votes. Nationalist support is the sum of Support for Batasuna, PNV, EA, PNV-EA, and Aralar, as a fraction of total votes. In all regressions, 2007 is excluded from the sample. All regressions control for a cubic in municipality size.

	Federal Left	Federal Right	Other local parties
	(1)	(2)	(3)
$I(\text{Post } 2007) \times I(\text{Ban in } 2007)$	0.00854 (0.00671)	-0.0109 (0.00534)	0.0756 (0.0296)
$I(\text{Post } 1995) \times I(\text{Ban in } 2007)$	0.00782 (0.00675)	0.00372 (0.00550)	-0.0312 (0.0210)
Municipality FE	✓	✓	✓
Pre-ban presence \times election FE	✓	✓	✓
N	1735	1735	1735
Municipalities	251	251	251

Standard errors clustered at the municipality level in parentheses. The placebo treatment is equal to zero in 1987, 1991, 1995, and equal to the treatment from 1999 onwards. Federal left is the sum of PSOE, IU, Podemos, Los Verdes, and Equo, as a fraction of total votes. Federal Right is the sum of PP, Ciudadanos, Unidad Alavesa, Centro Democrático y Social, and UPyD. Other local parties are votes that are neither for Nationalists, Federal Left or Federal Right. In all regressions, 2007 is excluded from the sample. All regressions control for a cubic in municipality size.

Table A3: Descriptive statistics, correlates of vote loss

Variable Name	Treated	Control	Total
$\ln(1+\# \text{ Terror Attacks})$	0.158 (0.487)	0.171 (0.440)	0.163 (0.468)
Close	0.204 (0.404)	0.153 (0.362)	0.184 (0.388)
$\ln(\text{Population in 2007})$	7.293 (1.737)	7.700 (1.461)	7.452 (1.644)
Batasuna Support in 2003	0.152 (0.152)	0.164 (0.147)	0.157 (0.150)
Municipalities	153	98	251

Treated: municipalities in which Batasuna was banned in 2007. Standard deviations in parentheses. Batasuna support is the sum of Batasuna votes and Null votes (in 2003, Null Votes). # Terror attacks is the number of ETA attacks in the municipality in the 10 years prior to the 2007 ban. Close is a dummy equal to one if in 2003, the absolute difference in vote share between Federal and Basque parties was below 10%. The source of information about the ETA attacks is the Global Terrorism Database.

Robustness check: regional elections - descriptive statistics and results

This section reports the effects of a longer ban in local elections on support for Batasuna in the Basque regional elections. This exercise is useful for two reasons. The first is that there was one regional election in 2005, just two years before the 2007 ban, which allows me to test whether treated and control municipalities were performing differently shortly before the treatment. The second reason is that Batasuna presented a common regional candidature across municipalities, which makes support for Batasuna in regional elections a better measure of preferences and support for the political movement. In addition to considerations of identity (Basque elections are seen as national elections for Basque Nationalists), Basque regional elections are high-stake elections, where Batasuna can obtain a significant share of the seats and wield influence, as opposed to the Spanish General elections and the European elections.²⁸

²⁸Batasuna did not always contest the Spanish elections. For instance, in 2000, in the last Spanish election before the 2003 ban, it asked voters not to participate.

At the regional level, Batasuna was banned for one election, but its legal status remained the same across all municipalities.

To estimate this effect, I use data on the municipal results of regional elections in the Basque Country from 1986 to 2012, which present a very similar party system to that of the local elections.²⁹ At the regional level, Batasuna escaped the ban in 2005 and so could contest that regional election, but it could not do so in 2009, when it called for a null vote, as it did in the local elections during the ban. Hence, I measure support in 2009 with the share of null votes, and to be consistent across years I define support for Batasuna as $\frac{\text{Batasuna votes} + \text{Null votes}}{\text{Total votes}}$.

I compare differences in support before (1986-2005) and after the heterogeneous enforcement of the ban in local elections (2009-2012), according to the legal status of Batasuna in the 2007 local election. The main summary statistics for either group are displayed in table A4. Table A5 reports difference-in-differences estimates. Column 1 in table A5 reports that a longer ban at the local level leads to less support in regional elections. The effect is small compared to local elections, where Batasuna is usually more successful, but still meaningful and significant. A longer ban reduced the support for Batasuna by 1.2pp, which is slightly less than 5% of their post-ban average support, and 7% of its standard deviation. In column 2, I add a lead of the treatment variable to see whether treatment and control municipalities were performing differently in the 2005 regional elections, just two years before the 2007 ban. The point estimate shows that this was not the case, supporting the common trends assumption.

The subsequent columns account for the decrease in vote for Batasuna under the ban, revealing a very similar pattern to that presented by the local elections: the point estimate falls from -0.0115 to 0.00144 after accounting for differences in voting behavior across treated and control municipalities in the 2007 local election.

²⁹The electoral districts are the three provinces making up the Basque Autonomous Community: Araba/Álava, Bizkaia/Vizcaya and Gipuzkoa/Guipúzcoa, and the electoral rule is d'Hondt within each district. Regional elections took place in 1986, 1990, 1994, 1998, 2001, 2005, 2009 and 2012.

Table A4: Descriptive statistics, regional elections

Variable Name	Treated		Control		Total	
	Pre	Post	Pre	Post	Pre	Post
Turnout	0.735 (0.0923)	0.699 (0.0743)	0.736 (0.0894)	0.693 (0.0705)	0.736 (0.0911)	0.697 (0.0728)
Batasuna votes	0.205 (0.134)	0.176 (0.210)	0.248 (0.130)	0.204 (0.228)	0.222 (0.134)	0.187 (0.217)
Batasuna Support	0.205 (0.134)	0.250 (0.174)	0.248 (0.130)	0.299 (0.171)	0.222 (0.134)	0.269 (0.174)
Null votes	0.00538 (0.00897)	0.0786 (0.108)	0.00507 (0.00510)	0.0989 (0.125)	0.00526 (0.00768)	0.0865 (0.116)
Population size	8238.0 (31561.8)	7965.3 (29896.8)	5173.3 (10771.1)	5171.8 (10351.1)	7034.1 (25540.8)	6874.6 (24242.7)
Municipalities	152	153	98	98	250	251
N	898	306	581	196	1479	502

Treated: municipalities in which Batasuna was banned in 2007. Batasuna support is the sum of Batasuna votes and Null votes. Standard deviations in parentheses.

Table A5: Effects of the 2007 ban on support for Batasuna in regional elections

	Batasuna Support			
	(1)	(2)	(3)	(4)
$I(\text{Post } 2007) \times I(\text{Ban in } 2007)$	-0.0115 (0.00583)	-0.0125 (0.00428)	-0.0125 (0.00428)	0.000398 (0.00581)
$I(\text{Post } 2005) \times I(\text{Ban in } 2007)$		0.00113 (0.00453)	0.00111 (0.00453)	0.00109 (0.00453)
Municipality FE	✓	✓	✓	✓
Province \times election FE	✓	✓	✓	✓
Pre-ban presence \times election FE	✓	✓	✓	✓
$I(\text{Post } 2007) \times \text{Pre-ban Support (local)}$			✓	✓
$I(\text{Post } 2007) \times \text{Support in } 2007 \text{ (local)}$				✓
N	1981	1981	1979	1979
Municipalities	251	251	250	250

Standard errors clustered at the municipality level in parentheses. In 2009, the vote share of Batasuna is computed as the share of null votes. Pre-ban support and support in 2007 refer to local elections. All regressions control for a cubic in municipality size.

Descriptive statistics and placebo tests, street violence

Table A6: Descriptive statistics, street violence, monthly

Variable Name	Treated		Control		Total	
	Pre	Post	Pre	Post	Pre	Post
$I(\text{Street Terrorist Episode})$	0.0297 (0.170)	0.0401 (0.196)	0.0290 (0.168)	0.0466 (0.211)	0.0294 (0.169)	0.0427 (0.202)
# Street Terrorist Episodes	0.0558 (0.428)	0.0705 (0.463)	0.0361 (0.232)	0.0583 (0.297)	0.0481 (0.365)	0.0657 (0.406)
Municipalities	153	153	98	98	251	251
N	1989	2142	1274	1372	3263	3514

Treated: municipalities in which Batasuna was banned in 2007. The unit of observation is a month by municipality. Standard deviations in parentheses.

Table A7: Effects of the 2007 ban on street violence, placebo test

	ST Episodes		
	(1)	(2)	(3)
$I(\text{Post May 2007}) \times I(\text{Ban in 2007})$	0.204 (0.0813)	0.000162 (0.0197)	-0.0155 (0.0196)
$I(\text{Post October 2006}) \times I(\text{Ban in 2007})$	-0.00966 (0.0191)	-0.00966 (0.0191)	-0.00966 (0.0191)
Sample	Until May 2007	Full	Except May 2007
Municipality FE	✓	✓	✓
Pre-ban presence \times Month FE	✓	✓	✓
N	3500	6750	6500
Municipalities	250	250	250

Standard errors clustered at the municipality level in parentheses. The dependent variable is given by the monthly number of street terrorism attacks.

	I(ST Episode)		
	(1)	(2)	(3)
$I(\text{Post May 2007}) \times I(\text{Ban in 2007})$	0.113 (0.0390)	0.00277 (0.0119)	-0.00574 (0.0119)
$I(\text{Post October 2006}) \times I(\text{Ban in 2007})$	-0.0163 (0.0126)	-0.0163 (0.0126)	-0.0163 (0.0126)
Sample	Until May 2007	Full	Except May 2007
Municipality FE	✓	✓	✓
Pre-ban presence \times Month FE	✓	✓	✓
N	3500	6750	6500
Municipalities	250	250	250

Standard errors clustered at the municipality level in parentheses. The dependent variable is a dummy equal to one if there is one or more episode of street terrorism in a month.

Additional robustness tests and heterogeneity analyses

Table A8: Effect on null votes, 2003 ban

	Null Votes (1)	Null Votes (2)	Batasuna Support (3)
$I(\text{Year}=2003) \times I(\text{Ban in 2007})$	-0.0110 (0.0191)	0.0202 (0.0188)	0.00597 (0.0163)
$I(\text{Post 2007}) \times I(\text{Ban in 2007})$			-0.0676 (0.0206)
Municipality FE	✓	✓	✓
Election FE	✓		
Pre-ban presence \times election FE		✓	✓
Sample	Until 2003	Until 2003	Full
N	1234	1234	1735
Municipalities	250	250	251

Standard errors clustered at the municipality level in parentheses. In all regressions, 2007 is excluded from the sample. All regressions control for a cubic in municipality size.

Table A9: Entropy Balancing estimates

	Batasuna Support (1)	All Basque Nationalists (2)
$I(\text{Post 2007}) \times I(\text{Ban in 2007})$	-0.0683 (0.0280)	-0.0625 (0.0283)
Municipality FE	✓	✓
Pre-ban presence \times election FE	✓	✓
N	1659	1659
Municipalities	237	237

Standard errors clustered at the municipality level in parentheses. Support for Batasuna is measured as the sum of Batasuna votes and Null votes, as a fraction of total votes. Nationalist support is the sum of Support for Batasuna, PNV, EA, PNV-EA, and Aralar, as a fraction of total votes. In all regressions, 2007 is excluded from the sample. All regressions control for a cubic in municipality size. Estimates weighted according to entropy balance weights, as described in the text.

Table A10: Allowing for heterogeneous post-ban effects

	Batasuna Support			
	(1)	(2)	(3)	(4)
$I(\text{Post } 2007) \times I(\text{Ban in } 2007)$	-0.0630 (0.0188)	-0.0698 (0.0176)	-0.0694 (0.0177)	-0.0702 (0.0185)
Municipality FE	✓	✓	✓	✓
Pre-ban presence \times election FE	✓	✓	✓	✓
Pre-ban characteristics	Population, Terror, and	Nationalists, Batasuna, and	All	All
\times Post 2007	Street Violence attacks	Spanish parties vote shares		+ pairwise interactions
N	1733	1733	1733	1733
Municipalities	250	250	250	250

Standard errors clustered at the municipality level in parentheses. Support for Batasuna is measured as the sum of Batasuna votes and Null votes, as a fraction of total votes. Population is measured with the average municipality size before the ban. Terror attacks are measured with the log of terror attacks from 1997 to 2006, Street Violence attacks are measured with the log of street violence attacks from April 2006 to April 2007. All political variable controls are pre-2003 averages. Nationalist support is the sum of Support for Batasuna, PNV, EA, PNV-EA, and Aralar, as a fraction of total votes. Spanish support is the sum of support for PP, CDS, UA, PSOE, PSOE, and IU, as a fraction of total votes. In all regressions, 2007 is excluded from the sample. All regressions control for a cubic in municipality size. Estimates weighted according to entropy balance weights, as described in the text.

Table A11: Heterogeneous effects by pre-ban support

	Batasuna Support	
	(1)	(2)
$I(\text{Year} \geq 2007) \times I(\text{Ban in } 2007)$	-0.0641 (0.0182)	-0.0777 (0.0185)
$I(\text{Year} \geq 2007) \times I(\text{Ban in } 2007) \times \text{Pre-Ban Support}$	0.0649 (0.177)	0.203 (0.194)
$I(\text{Year} \geq 2007) \times I(\text{Ban in } 2007) \times \text{Variance of Pre-Ban Support}$		-2.433 (0.592)
Municipality FE	✓	✓
Pre-ban presence \times election FE	✓	✓
$I(\text{Year} \geq 2007) \times \text{Pre-Ban Support}$	✓	✓
$I(\text{Year} \geq 2007) \times \text{Variance of Pre-Ban Support}$		✓
N	1476	1476
Municipalities	249	249

Standard errors clustered at the municipality level in parentheses. Support for Batasuna is measured as the sum of Batasuna votes and Null votes. Pre-ban support is a 1987-1999 average. All interacted variables are in deviations with respect to the cross-sectional mean. Range of pre-ban support: [-0.215,0.532]. Range of (de-meaned) variance of pre-ban support: [-0.0102,0.2626]. All regressions control for a cubic in municipality size. Sample: excluding 2003 and 2007.

Appendix B

Instrumental variables estimates

The heterogeneous enforcement of the ban in 2007 is related to the ability of the local party to find candidates without links with Batasuna in the past. Although the criteria employed by the public prosecutor for denouncing local lists was only known ex-post (i.e. after presenting the lists), Batasuna would have been able to anticipate these arguments. Indeed, their strategy was to deny any connection with EAE-ANV until the Courts had taken a final decision. If the ability to find *clean* candidates is related to the level of support for Batasuna, then this is absorbed by the municipality fixed effects, and D-i-D delivers the causal effect of interest. But if the differential ability is related to an unobservable change in support at the moment of the ban, then $E(\epsilon_{mt\text{Post } 2007} - \epsilon_{mt\text{Pre } 2007} | \text{Ban in 2007}) < E(\epsilon_{mt\text{Post } 2007} - \epsilon_{mt\text{Pre } 2007} | \text{Legal in 2007})$ and D-i-D would overestimate the negative effects of the ban on political support. To see whether this is the case, I use the rule established by the public prosecutor to denounce the Batasuna (EAE-ANV) lists to construct an instrument that relies on a weaker identification assumption.

In May 2007, the public prosecutor, drawing on police reports about links between the EAE-ANV lists and the previously banned organizations, sent a report to the Supreme Court aimed at preventing a number of EAE-ANV candidates from standing at that month's local election. In the preamble, the public prosecutor states his intention of denouncing lists when:

*“... from a **quantitative** point of view, at least three candidates with a relevant and direct relation with the banned party, including candidates in recent elections or party officials. To reach this figure, in principle, the elections before 1995 are not taken into account, nor the candidates known to belong to EAE-ANV”*

*“from a **qualitative** point of view, that any of the candidates participating in previous elections in direct representation of the banned parties had been elected and have had an institutional role (mayor, councilor, regional MP, provincial MP), which in itself is doubly indicative of their occupying a relevant place on the electoral lists; and, therefore, of their importance in the structure of the banned organization, and that is socially identified as a party official of the banned party, which grants the whole candidature a continuity that*

frustrates the execution of the verdict dictated by the Supreme Court”

In short, the public prosecutor denounced all lists with three or more former Batasuna candidates, among whom at least one had been elected as a councilor or MP for Batasuna in the past. Although the public prosecutor did not strictly adhere to this rule and not all lists denounced by the public prosecutor ended up being banned, this criterion based on observables and set ex-post (i.e. after the lists were presented) is a good predictor of decision subsequently taken by the Supreme Court, which gave rise to the actual heterogeneity in the legal status of Batasuna in the 2007 local election.³⁰ The definition of links used by the public prosecutor refers to candidates in the 1995 and 1999 local elections, considering the candidates on the lists of the banned candidatures in 2003 as merely minor links, although the latter had also received media attention before the prosecutor issued this report.³¹ Note also that the measure of such links simply requires a comparison of names across electoral lists and the results for different years, and, hence, it is perfectly observable and does not depend on police investigations.

I collect data on the quantity and quality of the links of each municipal list with previously banned organizations from two sources. The first corresponds to the report made by the public prosecutor, which contains information about the lists denounced.³² The second is a media report, which also contains information for the remaining candidatures.³³ I combine data from the two sources, taking the largest number of links as the true value for those municipalities for which information is available in both sources. For those municipalities for which there is no information available and where Batasuna was legal in 2007, I assume there were no links at all. For those municipalities for which there is no information available and which were not present in 2007, I assume that the number of links is equal to the size of the electoral list. Figure 11 shows the evolution over time of Batasuna’s electoral presence in four groups of municipalities, according to their number of former councilors (FCo) and former candidates (FCa). It shows that the group with three or more former candidates and one or more former councilors is the group that that was hit hardest by the 2007 ban, which

³⁰In addition to the public prosecutor, the State Lawyer also presented her own report to the Supreme Court, which largely overlaps with that of the public prosecutor, but without stating any precise criteria.

³¹El País, April 2007.

³²Link to the formal denunciation, accessed April 13, 2016.

³³Link to media information, accessed April 13, 2016.

is consistent with the criteria stated by the public prosecutor. Figure 12 shows the evolution over time of the electoral support for Batasuna for the same groups of municipalities. The figure shows how the group that was most affected by the ban -with three or more former candidates and one or more former councilors- was typically the third group with most support before the ban, but it became the group with the lowest level of support afterwards, starting in 2007 and persisting after the end of the ban. Hence, this is both the group that was hit hardest by the ban and whose relative performance was worse on the lifting of the ban.

Figure 11: Electoral presence, by number of links

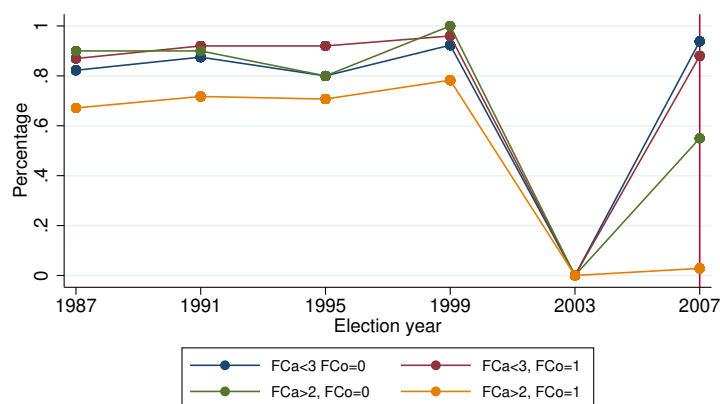
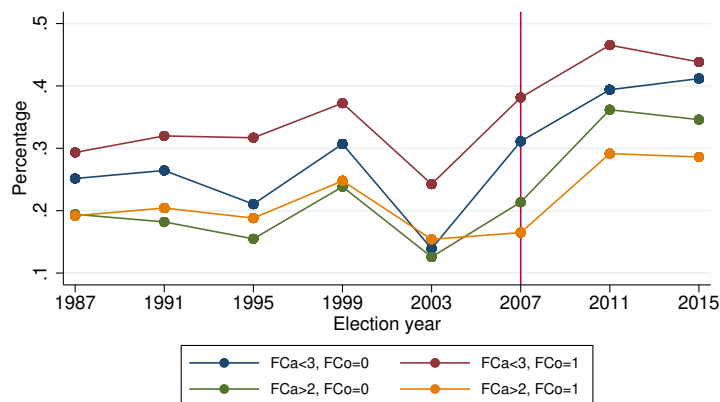


Figure 12: Electoral support, by number of links



I exploit this criteria, based on observables, to construct an instrument for Batasuna being legal in 2007, as follows. To simplify notation, I write the model in first differences, where Δ is the difference between the average post-ban level of support and the average pre-ban level of support. The reduced form reads as follows:

$$\Delta Support_m = \gamma I(FormerCouncilors_m \geq 1) + \pi I(FormerCandidates_m \geq 3) \\ + \theta(I(FormerCouncilors_m \geq 1) \times I(FormerCandidates_m \geq 3)) + \Delta \epsilon_m$$

Because the public prosecutor denounces lists with (a) at least one former councilor and (b) three or more former candidates, the excluded instrument is the interaction of (a) and (b). The IV is a triple difference estimator (DiDiD) given by:

$$\hat{\theta}^{DiDiD} = [\Delta \overline{Supp}_{FCa \geq 3}^{FCo \geq 1} - \Delta \overline{Supp}_{FCa < 3}^{FCo \geq 1}] - [\Delta \overline{Supp}_{FCa \geq 3}^{FCo=0} - \Delta \overline{Supp}_{FCa < 3}^{FCo=0}]$$

where *Supp* stands for support for Batasuna, *FCo* stands for Former Councilors and *FCa* stands for Former Candidates.

Its expectation is given by:

$$E(\theta_m | FCo \geq 1, FCa \geq 3) + [E(\Delta \epsilon_{mFCa \geq 3} - \Delta \epsilon_{mFCa < 3} | FCo \geq 1)] - [E(\Delta \epsilon_{mFCa \geq 3} - \Delta \epsilon_{mFCa < 3} | FCo = 0)]$$

Therefore, the identification assumption is that:

$$E[\Delta \epsilon_m | FCa_m \geq 3, FCo_m \geq 1] - E[\Delta \epsilon_m | FCa_m < 3, FCo_m \geq 1] \\ = E[\Delta \epsilon_m | FCa_m \geq 3, FCo_m = 0] - E[\Delta \epsilon_m | FCa_m < 3, FCo_m = 0]$$

Namely, that the pre-post difference in unobservables between those with more and less than three former candidates is the same for those with a former councilor and those without a former councilor. This is a parallel trend assumption for the change in the outcome (rather than for the outcome), over the number of former candidates (rather than over time), between those with a former councilor and those without. Table A12 reports descriptive statistics for each of these variables, by 2007 legal status.

Table A12: Descriptive statistics, linked candidates

Variable Name	Treated	Control	Total
I(Former Councilor)	0.915 (0.280)	0.265 (0.444)	0.661 (0.474)
Former Candidates	5.229 (1.819)	1.316 (1.118)	3.701 (2.481)
I(Former Candidates ≤ 3)	0.954 (0.210)	0.153 (0.362)	0.641 (0.481)
I(Former Councilor) \times I(Former Candidates ≤ 3)	0.895 (0.307)	0.0408 (0.199)	0.562 (0.497)
Municipalities	153	98	251

Treated: municipalities in which Batasuna was banned in 2007. For banned municipalities with no information, I(Former Councilor)=1 and Former Candidates=List size. Standard deviations in parentheses.

In practice, instead of first-differences I use fixed effects and interact all the relevant variables with a post-2007 dummy. I also include fixed effects (interacted with a post 2007 dummy) for the number of former candidates, so that the comparison is within municipalities with the same number of former candidates. Table A13 reports the first stage, which indicates that the interaction between more than three former candidates and a former councilor (i.e., the excluded instrument) predicts well the legal status in 2007. Table A14 reports instrumental variables results. Column 1 reports the baseline results, corresponding to the specification in table A13. Although the results are imprecise, the point estimate is qualitatively similar to the difference-in-differences estimate in table 1, albeit slightly larger in magnitude (8pp vs. 6pp). In column 2, I separately control for the number of former candidates for those with at least one former councilor and for those without. This is analogous to including group-specific trends in a standard difference-in-differences regression, and the results, which are similar, suggest that the identification assumption is plausible. In column 3, I exclude from the sample those municipalities for which I had no information on links and were not present in 2007 (and for which I assumed that the number of links was equal to the size of the electoral list); in column 4, I consider only municipalities with two or three former candidates. Overall, these estimates, although imprecise, suggest that the difference-in-differences estimates are not biased upwards and are capturing a causal effect of the total effect of the 2007 ban on electoral support.

Table A13: First Stage

	(1) $I(\text{Post 2007}) \times I(\text{Ban in 2007})$
$I(FCo \geq 1) \times I(FCa \geq 3) \times I(\text{Post 2007})$	0.473 (0.134)
Municipality FE	✓
Year FE	✓
Former Candidates FE $\times I(\text{Post 2007})$	✓
$I(FCo \geq 1) \times I(\text{Post 2007})$	✓
Kleiberger-Paap F	12.454
N	1735
Municipalities	251

Standard errors clustered at the municipality level in parentheses. 2007 is excluded from the sample. All regressions control for a cubic in municipality size.

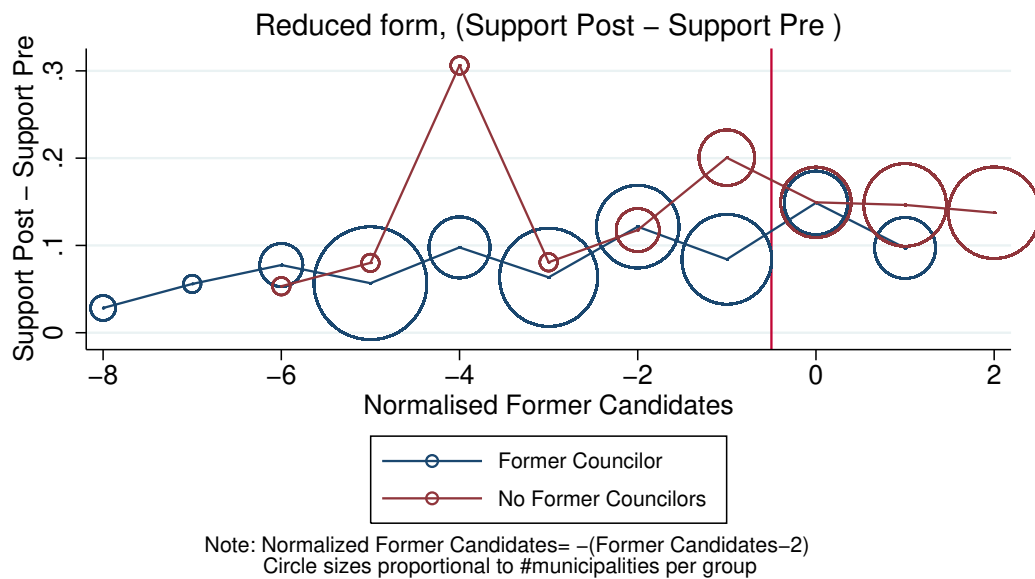
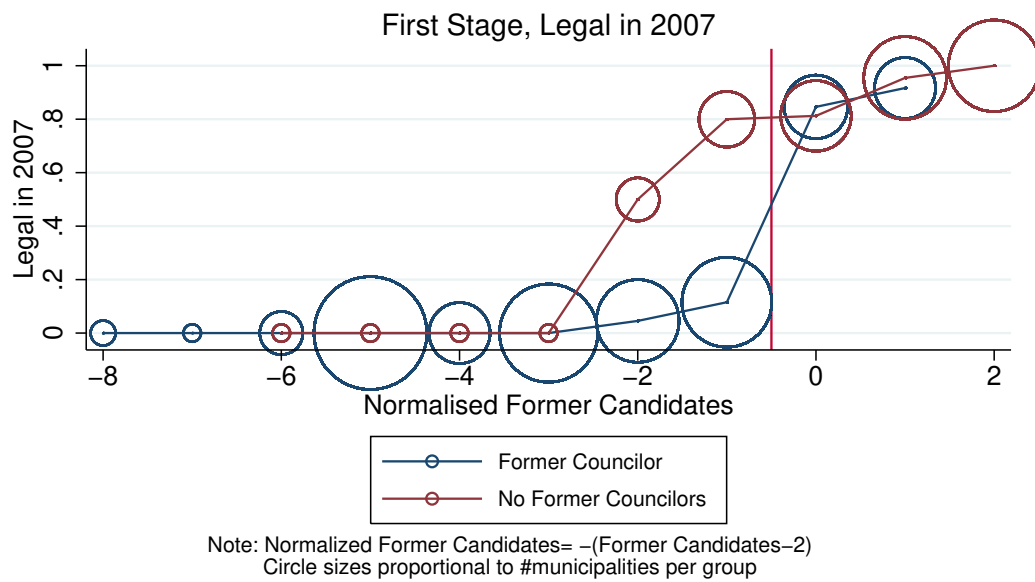
Table A14: Instrumental Variables estimates, support in local elections

	(1) Batasuna Support	(2) Batasuna Support	(3) Batasuna Support	(4) Batasuna Support
$I(\text{Post 2007}) \times I(\text{Ban in 2007})$	-0.0883 (0.0909)	-0.106 (0.0638)	-0.107 (0.0643)	-0.160 (0.0789)
Municipality FE	✓	✓	✓	✓
Year FE	✓	✓	✓	✓
Former Candidates FE $\times I(\text{Post 2007})$	✓	✓	✓	✓
$I(FCo \geq 1) \times I(\text{Post 2007})$	✓	✓	✓	✓
$FCa \times I(FCo \geq 1) \times I(\text{Post 2007})$		✓	✓	
Excluding missing links			✓	
Only $FCa = 2$ or $FCa = 3$				✓
Kleiberger-Paap F	12.454	26.069	25.745	19.018
N	1735	1735	1412	450
Municipalities	251	251	203	65

Standard errors clustered at the municipality level in parentheses. Support for Batasuna is measured as the sum of Batasuna votes and Null votes. In all regressions, 2007 is excluded from the sample. All regressions control for a cubic in municipality size.

Figure 13 displays the first stage and the reduced form. Circle sizes are proportional to the number of municipalities in each group, which is quite heterogeneous. In the figure, former candidates are normalized (Normalised Former Candidates = $-(\text{Former Candidates} - 2)$). Except for one outlier (corresponding to one municipality) at -4 normalized former candidates, which column 4 in table A14 shows is not driving the results, most of the action comes from the difference in legal status for those with 3 or 4 former candidates (-1 or -2 normalized candidates in the graph). It is within this group that the differences in legal status due to having a former councilor emerge, and it is within this group (in particular, for those with 3 former candidates, or -1 normalized former candidates in the figure) that I observe the differences in support driving the results.

Figure 13: Effect of the 2007 ban on support for Batasuna, local electons, IV



Regarding the effect on support in regional elections, table A15 reports instrumental variables estimates. Column 1 reports the baseline results, corresponding to the first stage specification in table A13. Although the results are imprecise, the point estimate is qualitatively similar to the difference-in-differences estimate in column 1 of table A5. In column 2, I separately control for the number of former candidates for those with at least one former councilor and for those without, which is analogous to including group-specific trends in a standard difference-in-differences regression. In column 3, I exclude from the sample those municipalities for which I had no information on links and were not present in 2007 (and for which I assumed that the number of links was equal to the size of the electoral list); in column 4, I consider only municipalities with two or three former candidates. The results are quite stable across specifications, with point estimates ranging between 1pp and 2pp.

Table A15: Instrumental Variables estimates, support in regional elections

	(1)	(2)	(3)	(4)
	Batasuna Support	Batasuna Support	Batasuna Support	Batasuna Support
$I(\text{Post 2007}) \times I(\text{Ban in 2007})$	-0.0216 (0.0307)	-0.0182 (0.0227)	-0.0114 (0.0223)	-0.0150 (0.0281)
Municipality FE	✓	✓	✓	✓
Year FE	✓	✓	✓	✓
Former Candidates FE $\times I(\text{Post 2007})$	✓	✓	✓	✓
$I(FCo \geq 1) \times I(\text{Post 2007})$	✓	✓	✓	✓
$FCa \times I(FCo \geq 1) \times I(\text{Post 2007})$		✓	✓	
Excluding missing links			✓	
Only $FCa = 2$ or $FCa = 3$				✓
Kleibergen-Paap F	11.425	21.182	20.724	14.416
N	1981	1981	1609	514
Municipalities	251	251	203	65

Standard errors clustered at the municipality level in parentheses. All regressions control for a cubic in municipality size.